

Solving Optimization Problems with Diseconomies of Scale via Decoupling

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Abstract—We present a new framework for solving optimization problems with a diseconomy of scale. In such problems, our goal is to minimize the cost of resources used to perform a certain task. The cost of resources grows superlinearly, as x^q , $q \geq 1$, with the amount x of resources used. We define a novel linear programming relaxation for such problems, and then show that the integrality gap of the relaxation is A_q , where A_q is the q -th moment of the Poisson random variable with parameter 1. Using our framework, we obtain approximation algorithms for the Minimum Energy Efficient Routing, Minimum Degree Balanced Spanning Tree, Load Balancing on Unrelated Parallel Machines, and Unrelated Parallel Machine Scheduling with Nonlinear Functions of Completion Times problems.

Our analysis relies on the decoupling inequality for nonnegative random variables. The inequality states that

$$\left\| \sum_{i=1}^n X_i \right\|_q \leq C_q \left\| \sum_{i=1}^n Y_i \right\|_q,$$

where X_i are independent nonnegative random variables, Y_i are possibly dependent nonnegative random variable, and each Y_i has the same distribution as X_i . The inequality was proved by de la Peña in 1990. However, the optimal constant C_q was not known. We show that the optimal constant is $C_q = A_q^{1/q}$.

I. INTRODUCTION

In this paper, we study combinatorial optimization problems with a diseconomy of scale. We consider problems in which we need to minimize the cost of resources used to accomplish a certain task. Often, the cost grows linearly with the amount of resources used. In some applications, the cost is sublinear e.g., if we can get a discount when we buy resources in bulk. Such phenomenon is known as “economy of scale”. However, in many applications the cost is superlinear. In such cases, we say that the cost function exhibits a “diseconomy of scale”. A good example of a diseconomy of scale is the cost of energy used for computing. Modern hardware can run at different processing speeds. As we increase the speed, the energy consumption grows superlinearly. It can be modeled as a function $P(s) = cs^q$ of the processing speed s , where c and q are parameters that depend on the specific hardware. Typically, $q \in (1, 3]$ (see e.g., [2], [20], [37]).

As a running example, consider the Minimum Power Routing problem studied by Andrews, Anta, Zhang, and Zhao [3]. We are given a graph $G = (V, E)$ and a set of demands $\mathcal{D} = \{(d_i, s_i, t_i)\}$. Our goal is to route d_i ($d_i \in \mathbb{N}$) units of demand i from the source $s_i \in V$ to the destination $t_i \in V$ such that every demand i is routed along a single path p_i (i.e.

we want to find an unsplittable multi-commodity flow). We want to minimize the energy cost. Every link (edge) $e \in E$ uses $f_e(x_e) = c_e x_e^q$ units of power, where c_e is a scaling parameter depending on the link e , and x_e is the load on e .

The straightforward approach to solving this problem is as follows. We define a mathematical programming relaxation that routes demands fractionally. It sends $y_{i,p} d_i$ units of demand via the path p connecting s_i to t_i . We require that $\sum_p y_{i,p} = 1$ for every demand i . The objective function is to minimize

$$\min \sum_{e \in E} c_e x_e^q = \min \sum_{e \in E} c_e \left(\sum_{p: e \in p} y_{i,p} d_i \right)^q,$$

where $x_e = \sum_{p: e \in p} y_{i,p} d_i$ is the load on the link e . This relaxation can be solved in polynomial time, since the objective function is convex (for $q \geq 1$). But, unfortunately, the integrality gap of this relaxation is $\Omega(n^{q-1})$ [3]. Andrews et al. [3] gave the following integrality gap example. Consider two vertices s and t connected via n disjoint paths. Our goal is to route 1 unit of flow integrally from s to t . The optimal solution pays 1. The LP may cheat by routing $1/n$ units of flow via n disjoint paths. Then, it pays only $n \times (1/n)^q = n^{1-q}$.

For the case of uniform demands, i.e., for the case when all $d_i = d$, Andrews et al. [3] suggested a different objective function:

$$\min \sum_{e \in E} c_e \max\{x_e^q, d^{q-1} x_e\}.$$

The objective function is valid, because in the integral case, x_e must be a multiple of d , and thus $x_e^q \geq d^{q-1} x_e$. Andrews et al. [3] proved that the integrality gap of this relaxation is a constant. Bampis et al. [9] improved the bound to the *fractional Bell number* A_q that is defined as follows: A_q is the q -th moment of the Poisson random variable P_1 with parameter 1 (see Figure 2 in Appendix A). I.e.,

$$A_q = \mathbb{E}[P_1^q] = \sum_{t=1}^{+\infty} t^q \frac{e^{-1}}{t!}. \quad (1)$$

For the case of general demands no constant approximation was known. The best known approximation due to Andrews et al. [3] was $O(k + \log^{q-1} \Delta)$ where $k = |\mathcal{D}|$ is the number of demands and $\Delta = \max_i d_i$ is the size of the largest demand (Theorem 8 in [3]).

In this work, we give an A_q -approximation algorithm for the general case and thus close the gap between the case of uniform and non-uniform demands. Our approximation

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algorithm uses a general framework for solving problems with a diseconomy of scale which we present in this paper. We use this framework to obtain approximation algorithms for several other combinatorial optimization problems. We give $A_q^{1/q}$ -approximation algorithm for Load Balancing on Unrelated Parallel Machines (see Section II-B), $2^q A_q$ -approximation algorithm for Unrelated Parallel Machine Scheduling with Non-linear Functions of Completion Times (see the full version of the paper) and A_q -approximation algorithm for the Minimum Degree Balanced Spanning Tree problem (see Section II-C). The best previously known bound for the first problem with $q \in [1, 2]$ was $2^{1/q}$ (see Figure 3 for comparison). The bound is due to Kumar, Marathe, Parthasarathy and Srinivasan [22]. There were no known approximation guarantees for the latter problems.

In the analysis, we use the de la Peña decoupling inequality [26], [27].

Theorem I.1 (de la Peña [26], [27]). *Let Y_1, \dots, Y_n be jointly distributed nonnegative (non-independent) random variables, and let X_1, \dots, X_n be independent random variables such that each X_i has the same distribution as Y_i . Then, for every $q \geq 1$,*

$$\left\| \sum_{i=1}^n X_i \right\|_q \leq C_q \left\| \sum_{i=1}^n Y_i \right\|_q, \quad (2)$$

for some universal constant C_q .

The optimal value of the constant C_q was not known. The original proof of de la Peña relies on more general inequalities and does not give specific constants. We give a direct proof of this inequality. We show that the inequality holds for $C_q = A_q^{1/q}$ and moreover this bound is tight.

Theorem I.2. *Inequality (2) holds for $C_q = A_q^{1/q}$, where A_q is the fractional Bell number (see Equation (1) and Figure 2) Moreover, $A_q^{1/q}$ is the optimal upper bound on C_q .*

In Section VI (see Corollary VI.2), we extend this theorem to negatively associated random variables X_i .

A. General Framework

We now describe the general framework for solving problems with a diseconomy of scale. We consider optimization problems with n decision variables $y_1, \dots, y_n \in \{0, 1\}$. We assume that the objective function equals the sum of k terms, where the j -th term is of the form

$$c_j \left(\sum_{i=1}^n d_{ij} y_i \right)^{q_j},$$

here $c_j \geq 0$, $d_{ij} \geq 0$ and $q_j \geq 1$ are parameters. The vector $y = (y_1, \dots, y_n)$ must satisfy the constraint $y \in \mathcal{P}$ for some polytope $\mathcal{P} \subset [0, 1]^n$. Therefore, the optimization problem can

be written as the following boolean convex program (IP):

$$\min \sum_{j \in [k]} c_j \left(\sum_{i \in [n]} d_{ij} y_i \right)^{q_j} \quad (3)$$

$$y \in \mathcal{P} \quad (4)$$

$$y \in \{0, 1\}^n \quad (5)$$

We assume that we can optimize any linear function over the polytope \mathcal{P} in polynomial time (e.g., \mathcal{P} is defined by polynomially many linear inequalities, or there exists a separation oracle for \mathcal{P}). Thus, if we replace the integrality constraint (5) with the relaxed constraint $y \in [0, 1]^n$ (which is redundant, since $\mathcal{P} \subset [0, 1]^n$), we will get a convex programming problem that can be solved in polynomial time (see [11]). However, as we have seen in the example of Minimum Power Routing, the integrality gap of the relaxation can be as large as $\Omega(n^{q-1})$.

In this work, we introduce a linear programming relaxation of (3)-(5) that has an integrality gap of A_q (where $q = \max_j q_j$) under certain assumptions on the polytope \mathcal{P} . We define auxiliary variables z_{jS} for all $S \subset [n]$ and $j \in [k]$. In the integral solution, $z_{jS} = 1$ if and only if $y_i = 1$ for $i \in S$ and $y_i = 0$ for $i \notin S$.

$$\min \sum_{j \in [k]} \sum_{S \subseteq [n]} c_j \left(\sum_{i \in S} d_{ij} \right)^{q_j} z_{jS} \quad (6)$$

$$y \in \mathcal{P}, \quad (7)$$

$$\sum_{S \subseteq [n]} z_{jS} = 1, \quad \forall j \in [k] \quad (8)$$

$$\sum_{S: i \in S} z_{jS} = y_i, \quad \forall i \in [n], j \in [k] \quad (9)$$

$$z_{jS} \geq 0, \quad \forall S \subseteq [n], j \in [k] \quad (10)$$

Remark I.3. *In the integral solution, $z_{j'S} = z_{j''S}$ for all j' and j'' . The reason why we introduced many copies of the same integral variable z_S to the LP is that the LP above is easier to solve than the LP with an extra constraint $z_{j'S} = z_{j''S}$.*

Optimization problem (6)-(10) is a relaxation of the original problem (3)-(5). The LP has exponentially many variables. We show, however, that the optimal solution to this LP can be found in polynomial time up to an arbitrary accuracy $(1 + \varepsilon)$. We say that y is a $(1 + \varepsilon)$ -approximately optimal solution if the cost of the solution is at most $(1 + \varepsilon)OPT$, where OPT is the cost of the optimal solution.

Theorem I.4. *Suppose that there exists a polynomial time separation oracle for the polytope \mathcal{P} . Then, for every ε and q , there exists a polynomial time algorithm that finds a $(1 + \varepsilon)$ -approximately optimal solution to LP (6)-(10).*

We then prove the following theorem.

Theorem I.5. *Let $D_j = \{i : d_{ij} \neq 0\}$. Assume that there exists a randomized algorithm R that given a $y \in \mathcal{P}$, returns a random integral point $R(y)$ in $\mathcal{P} \cap \{0, 1\}^n$ such that*

- 1) $\Pr(R_i(y) = 1) = y_i$ for all i (where $R_i(y)$ is the i -th coordinate of $R(y)$);
- 2) Random variables $\{R_i(y)\}_{i \in \mathcal{D}_j}$ are independent or negatively associated (see Section VI for the definition) for every j .

Then, for every feasible solution (y^*, z^*) to LP (6)-(10), we have

$$\mathbb{E} \left[\sum_{j \in [k]} c_j \left(\sum_{i \in [n]} d_{ij} R_i(y^*) \right)^{q_j} \right] \leq A_q \sum_{j \in [k]} \sum_{S \subseteq [n]} c_j \left(\sum_{i \in S} d_{ij} \right)^{q_j} z_{jS}^*, \quad (11)$$

where $q = \max_j q_j$ and A_q is the fractional Bell number (see (1)). Particularly, since LP (6)-(10) is a relaxation for IP (3)-(5), if (y^*, z^*) is a $(1 + \varepsilon)$ -approximately optimal solution to LP (6)-(10), then

$$\mathbb{E} \left[\sum_{j \in [k]} c_j \left(\sum_{i \in [n]} d_{ij} R_i(y^*) \right)^{q_j} \right] \leq (1 + \varepsilon) A_q IP,$$

where IP is the optimal cost of the boolean convex program (3)-(5).

This theorem guarantees that an algorithm R satisfying conditions (1) and (2) has an approximation ratio of $(1 + \varepsilon)A_q$.

In the next section, Section II, we show how to use the framework to obtain A_q approximation algorithms for four different combinatorial optimization problems. Then, in Section III, we give an efficient algorithm for solving LP (6)-(10). In Section IV, we prove the main theorem – Theorem I.5. The proof easily follows from the decoupling inequality, which we prove in Section V. Finally, in Section VII, we describe some generalizations of our framework.

II. APPLICATIONS

In this section, we show applications of our general technique. We start with the problem discussed in the introduction – Energy Efficient Routing. Recall, that Andrews et al. [3] gave an $O(k + \log^{q-1} \Delta)$ -approximation algorithm for this problem where $k = |\mathcal{D}|$ and $\Delta = \max_{i \in \mathcal{D}} d_i$ (Theorem 8 in [3]). We give an A_q -approximation algorithm.

A. Energy Efficient Routing

We write a standard integer program. Each variable $y_{i,e} \in \{0, 1\}$ indicates whether the edge e is used to route the flow from s_i to t_i . Below, $\Gamma^+(u)$ denotes the set of edges outgoing from u ; $\Gamma^-(u)$ denotes the set of edges incoming to u .

$$\min \sum_{e \in E} c_e \left(\sum_{i \in \mathcal{D}} d_i y_{i,e} \right)^{q_e} \quad (12)$$

$$\sum_{e \in \Gamma^+(u)} y_{i,e} = \sum_{e \in \Gamma^-(u)} y_{i,e}, \quad \forall i, u \in V \setminus \{s_i, t_i\} \quad (13)$$

$$\sum_{e \in \Gamma^+(s_i)} y_{i,e} = 1, \quad \forall i \quad (14)$$

$$\sum_{e \in \Gamma^-(t_i)} y_{i,e} = 1, \quad \forall i \quad (15)$$

$$y_{i,e} \in \{0, 1\}, \quad \forall i, e \in E \quad (16)$$

Using Theorem I.4, we obtain an almost optimal fractional solution (y, z) of LP relaxation (6)-(10) of IP (12)-(16). We apply randomized rounding in order to select a path for each demand. Specifically, for each demand $i \in \mathcal{D}$, we consider the standard flow decomposition into paths: In the decomposition, each path p connecting s_i to t_i has a weight $\lambda_{i,p} \in \mathbb{R}^+$. For every edge e , $\sum_{p: e \in p} \lambda_{i,p} = d_i y_{i,e}$; and $\sum_p \lambda_{i,p} = d_i$. For each i , the approximation algorithm picks one path p connecting s_i to t_i at random with probability $\lambda_{i,p}/d_i$, and routes all demands from s_i to p_i via p . Thus, the algorithm always obtains a feasible solution.

We verify that the integral solution corresponding to this combinatorial solution satisfies the conditions of Theorem I.5. Let $R_{i,e}(y)$ be the integral solution, i.e., let $R_{i,e}(y) = 1$ if the edge e is chosen in the path connecting s_i and t_i . First, $R_{i,e}(y) = 1$ if the path connecting s_i and t_i contains e , thus

$$\Pr(R_{i,e}(y) = 1) = \sum_{p: e \in p} \lambda_{i,p}/d_i = y_{i,e}.$$

Second, the paths for all demands are chosen independently. Each $R_{i,e}(y)$ depends only on paths that connect s_i to t_i . Thus all random variables $R_{i,e}(y)$ (for a fixed e) are independent. Therefore, by Theorem I.5, the cost of the solution obtained by the algorithm is bounded by $(1 + \varepsilon)A_q OPT$, where OPT is the cost of the optimal solution to the integer program which is exactly equivalent to the Minimum Energy Efficient Routing problem.

B. Load Balancing on Unrelated Parallel Machines

We are given n jobs and m machines. The processing time of the job $j \in [n]$ assigned to the machine $i \in [m]$ is $p_{ij} \geq 0$. The goal is to assign jobs to machines to minimize the ℓ_q -norm of machines loads. Formally, we partition the set of jobs into m sets S_1, \dots, S_m to minimize $\left(\sum_{i \in [m]} \left(\sum_{j \in S_i} p_{ij} \right)^q \right)^{1/q}$. This is a classical scheduling problem which is used to model load balancing in practice¹. It was previously studied by Azar and Epstein [6] and by Kumar, Marathe, Parthasarathy and Srinivasan [22]. Particular, for $q \in (1, 2]$ the best known approximation algorithm has performance guarantee $2^{1/q}$ [22] (Theorem 4.4). We give

¹A slight modification of the problem, where the objective is $\min \sum_{i \in [m]} \left(\sum_{j \in S_i} p_{ij} \right)^q$, can be used for energy efficient scheduling. Imagine that we need to assign n jobs to m processors/cores so that all jobs are completed by a certain deadline D . We can run processors at different speeds s_i . To meet the deadlines we must set $s_i = D^{-1} \sum_{j \in S_i} p_{ij}$. The total power consumption is proportional to $D \times \sum_{i=1}^m s_i^q = D^{1-q} \times \sum_{i=1}^m \left(\sum_{j \in S_i} p_{ij} \right)^q$. For this problem, our algorithm gives A_q approximation.

$\sqrt[q]{A_q}$ -approximation algorithm substantially improving upon previous results (see Figure 3).

We formulate the unrelated parallel machine scheduling problem as a boolean nonlinear program:

$$\min \sum_{i \in [m]} \left(\sum_{j \in [n]} p_{ij} x_{ij} \right)^q \quad (17)$$

$$\sum_{i \in [m]} x_{ij} = 1, \quad \forall j \in [n] \quad (18)$$

$$x_{ij} \in \{0, 1\}, \quad \forall i \in [m], j \in [n] \quad (19)$$

Using Theorem I.4, we obtain an almost optimal fractional solution (x, z) of the LP relaxation (6)–(10) corresponding to the IP (17)–(19). We use the straightforward randomized rounding: we assign each job j to machine i with probability x_{ij} . We claim that, by Theorem I.5, the expected cost of our integral solution is upper bounded by A_q times the value of the fractional solution (x, z) . Indeed, the probability that we assign a job j to machine i is exactly equal to x_{ij} ; and we assign job j to machine i independently of other jobs. That implies that our approximation algorithm has a performance guarantee of $\sqrt[q]{A_q}$ for the ℓ_q -norm objective.

C. Degree Balanced Spanning Tree Problem

We are given an undirected graph $G = (V, E)$ with edge weights $w_e \geq 0$. The goal is to find a spanning tree T minimizing the objective function

$$f(T) = \sum_{v \in V} \left(\sum_{e \in \delta(v) \cap T} w_e \right)^q, \quad (20)$$

where $\delta(v)$ is the set of edges in E incident to the vertex v . For $q = 2$, a more general problem was considered before in the Operations Research literature [5], [23], [25], [28] under the name of Adjacent Only Quadratic Spanning Tree Problem. A related problem, known as Degree Bounded Spanning Tree, received a lot of attention in Theoretical Computer Science [33], [16]. We are not aware of any previous work on Degree Balanced Spanning Tree Problem.

Let x_e be a boolean decision variable such that $x_e = 1$ if we choose edge $e \in E$ to be in our solution (tree) T . We formulate our problem as the following convex boolean optimization problem

$$\min \sum_{v \in V} \left(\sum_{e \in \delta(v)} w_e x_e \right)^q$$

$$x \in \mathcal{B}(\mathcal{M})$$

$$x_e \in \{0, 1\}, \quad \forall e \in E,$$

where $\mathcal{B}(\mathcal{M})$ is the base polymatroid polytope of the graphic matroid in graph G . We refer the reader to Schrijver's book [30] for the definition of the matroid. Using Theorem I.4, we obtain an almost optimal fractional solution x^* of LP relaxation (6)–(10) corresponding to the above integer problem.

Following Calinescu et al. [7], we define the continuous extension of the objective function (20) for any fractional solution x'

$$F(x') = \sum_{S \subseteq [n]} f(S) \prod_{e \in S} x'_e \prod_{e \notin S} (1 - x'_e),$$

i.e. $F(x')$ is equal to the expected value of the objective function (20) for the set of edges sampled independently at random with probabilities $x'_e, e \in E$. The function F cannot be computed exactly, but it can be approximated up to any factor $(1 + \varepsilon)$ via sampling. By Theorem I.5, we get the bound $F(x^*) \leq A_q \cdot LP^*$, where LP^* is the value of the LP relaxation (6)–(10) on the fractional solution x^* .

The rounding phase of the algorithm implements the pipage rounding technique [1] adopted to polymatroid polytopes by Calinescu et al. [7]. Calinescu et al. [7] showed that given a matroid \mathcal{M} and a fractional solution $x \in \mathcal{B}(\mathcal{M})$, one can efficiently find two elements, or two edges in our case, e' and e'' such that the new fractional solution $\tilde{x}(\varepsilon)$ defined as $\tilde{x}_{e'}(\varepsilon) = x_{e'} + \varepsilon$, $\tilde{x}_{e''}(\varepsilon) = x_{e''} - \varepsilon$ and $\tilde{x}_e(\varepsilon) = x_e$ for $e \notin \{e', e''\}$ is feasible in the base polymatroid polytope for small positive and for small negative values of ε .

They also showed that if the objective function $f(S)$ is submodular then the function of one variable $F(\tilde{x}(\varepsilon))$ is convex. In our case, the objective function $f(S)$ is supermodular which follows from a more general folklore statement.

Fact II.1. *The function $f(S) = g(\sum_{i \in S} w_i)$ is supermodular if $w_i \geq 0$ for $i \in [n]$ and $g(x)$ is a convex function of one variable.*

Therefore, the function $F(\tilde{x}(\varepsilon))$ is concave. Hence, we can apply the pipage rounding directly: We start with the fractional solution x^* . At every step, we pick e' and e'' (using the algorithm from [7]) and move to $\tilde{x}(\varepsilon)$ with $\varepsilon = \varepsilon_1 = -\min\{x_{e'}, 1 - x_{e''}\}$ or $\varepsilon = \varepsilon_2 = \min\{1 - x_{e'}, x_{e''}\}$ whichever minimizes the concave function $F(\tilde{x}(\varepsilon))$ on the interval $[\varepsilon_1, \varepsilon_2]$. We stop when the current solution \tilde{x} is integral.

At every step, we decrease the number of fractional variables x_e by at least 1. Thus, we terminate the algorithm in at most $|E|$ iterations. The value of the function $F(\tilde{x})$ never increases. So the cost of the final integral solution is at most the cost of the initial fractional solution x^* , which, in turn, is at most $A_q \cdot LP^*$.

Note, that we have not used any special properties of graphic matroids. The algorithm from [7] works for general matroids accessible through oracle calls. So we can apply our technique to more general problems where the objective is to minimize a function like (20) subject to base matroid constraints.

III. PROOF OF THEOREM I.4

We now give an efficient algorithm for finding $(1 + \varepsilon)$ approximately optimal solution to LP (6)–(10).

Proof of Theorem I.4: We first transform our instance to make all d_{ij} 's integral and polynomially bounded in nk/ε . This can be done using the standard scaling technique: round

down all d_{ij} to be multiples of $\varepsilon' = d\varepsilon/(3kqn)$, where $d = \max d_{ij}$. By doing so we may decrease the optimal value of the program by a factor of at most $(1+\varepsilon)$. Then, we rescale all d_{ij} by $1/\varepsilon'$ to make them integral. So, from now on, we will assume that all d_{ij} are integral and polynomially bounded.

Observe that for every $y \in \mathcal{P}$, there exists a z such that the pair (y, z) is a feasible solution to LP (6)-(10). For example, one such z is defined as $z_{jS} = \prod_{i \in S} y_i \prod_{i \notin S} (1 - y_i)$. Of course, this particular z may be suboptimal. However, it turns out, as we show below, that for every y , we can find the optimal z efficiently. Let us denote the minimal cost of the j -th term in (6) for a given $y \in \mathcal{P}$ by $F_j(y)$. That is, $F_j(y)$ is the cost of the following LP. The variables of the LP are z_{jS} . The parameters $y \in \mathcal{P}$ and $j \in [k]$ are fixed.

$$\min \sum_{S \subseteq [n]} \left(\sum_{i \in S} d_{ij} \right)^{q_j} z_{jS} \quad (21)$$

$$\sum_{S \subseteq [n]} z_{jS} = 1 \quad (22)$$

$$\sum_{S: i \in S} z_{jS} = y_i, \quad \forall i \in [n] \quad (23)$$

$$z_{jS} \geq 0, \quad \forall S \subseteq [n] \quad (24)$$

Now, LP (6)-(10) can be equivalently rewritten as (below y is the variable).

$$\min \sum_{j \in [k]} F_j(y) \quad (25)$$

$$y \in \mathcal{P} \quad (26)$$

The functions $F_j(y)$ are convex². In Lemma III.1 (see below), we prove that LP (21)-(24) can be solved in polynomial time, and thus the functions $F_j(y)$ can be computed efficiently. The algorithm for finding $F_j(y)$ also returns a subgradient of F_j at y . Hence, the minimum of convex problem (25)-(26) can be found using the ellipsoid method. Once the optimal y^* is found, we find z^* by solving LP (21)-(24) for y^* . ■

Lemma III.1. *There exists a polynomial time algorithm for computing F_j and finding a subgradient of F_j .*

Proof: We need to solve LP (21)-(24). We write the dual LP. We introduce a variable ξ for constraint (22) and variables η_i for constraints (23).

$$\max \quad \xi + \sum_i \eta_i y_i \quad (27)$$

$$\xi + \sum_{i \in S} \eta_i \leq \left(\sum_{i \in S} d_{ij} \right)^{q_j}, \quad \forall S \subseteq [n] \quad (28)$$

The LP has exponentially many constraints. However, finding a violated constraint is easy. To do so, we guess $B^* = \sum_{i \in S^*} d_{ij}$ for the set S^* violating the constraint. That is

²If z^* and z^{**} are the optimal solutions for vectors y^* and y^{**} , then $\lambda z^* + (1-\lambda)z^{**}$ is a feasible solution for $\lambda y^* + (1-\lambda)y^{**}$. Hence, $F_j(\lambda y^* + (1-\lambda)y^{**}) \leq \lambda F_j(y^*) + (1-\lambda)F_j(y^{**})$. See the full version of the paper for more details.

possible, since all d_{ij} are polynomially bounded, and so is B^* . Then we solve the maximum knapsack problem

$$\max_{S \subseteq [n]} \sum_{i \in S} \eta_i$$

$$\sum_{i \in S} d_{ij} = B^*$$

using the standard dynamic programming algorithm and obtain the optimal set S . The knapsack problem is polynomially solvable, since B^* is polynomially bounded. If $\xi + \sum_{i \in S} \eta_i \leq (B^*)^{q_j}$, then constraint (28) is violated for the set S ; otherwise all constraints (28) are satisfied.

Let (ξ^*, η^*) be the optimal solution of the dual LP. The value of the function $F_j(y)$ equals the objective value of the dual LP. A subgradient of F_j at y is given by the equation

$$\tilde{y} \mapsto \xi^* + \sum_i \eta_i^* \tilde{y}_i. \quad (29)$$

This is a subgradient of F_j , since (ξ^*, η^*) is a feasible solution of the dual LP for every \tilde{y} (note that constraint (28) does not depend on y), and, hence, (29) is a lower bound on $F_j(\tilde{y})$. ■

IV. PROOF OF THEOREM I.5

In this section, we prove the main theorem – Theorem I.5.

Proof of Theorem I.5: The theorem easily follows from the decoupling inequality (Theorem I.2 and Corollary VI.2). Consider a feasible solution (y^*, z^*) to IP (3)-(5). We prove inequality (11) term by term. That is, for every j we show that

$$\mathbb{E} \left[\left(\sum_{i \in D_j} d_{ij} R_i(y^*) \right)^{q_j} \right] \leq A_{q_j} \sum_{S \subseteq [n]} \left(\sum_{i \in S \cap D_j} d_{ij} \right)^{q_j} z_{jS}^*. \quad (30)$$

Recall that $D_j = \{i : d_{ij} \neq 0\}$. Above, we dropped terms with $i \notin D_j$, since if $i \notin D_j$, then $d_{ij} = 0$.

Fix a $j \in [n]$. Define random variables Y_i for $i \in D_j$ as follows: Pick a random set $S \subset [n]$ with probability z_{jS} , and let $Y_i = d_{ij}$ if $i \in S$, and $Y_i = 0$ otherwise. Note that random variables Y_i are dependent. We have

$$\Pr(Y_i = d_{ij}) = \sum_{S: i \in S} \Pr(S) = \sum_{S: i \in S} z_{jS}^* = y_i^*.$$

It is easy to see that

$$\mathbb{E} \left[\left(\sum_{i \in [n]} Y_i \right)^{q_j} \right] = \sum_{S \subseteq [n]} \left(\sum_{i \in S} d_{ij} \right)^{q_j} z_{jS}.$$

The right hand side is simply the definition of the expectation on the left hand side. Now, let $X_i = d_{ij} R_i(y^*)$ for $i \in D_j$. Note that by conditions of the theorem, $\Pr(X_i = d_{ij}) = y_i^* = \Pr(Y_i = d_{ij})$ (by condition (1)). Thus, each X_i has the same distribution as Y_i . Furthermore, X_i 's are independent or negatively associated (by condition (2)). Therefore, we can apply the decoupling inequality

$$\mathbb{E} \left[\left(\sum_{i \in D_j} X_i \right)^{q_j} \right] \leq A_{q_j} \mathbb{E} \left[\left(\sum_{i \in D_j} Y_i \right)^{q_j} \right].$$

The left hand side of the inequality equals the left hand side of (30), the right hand side of the inequality equals the right hand side of (30). Hence, inequality (30) holds. \blacksquare

V. DECOUPLING INEQUALITY

In this section, we prove the decoupling inequality (Theorem I.2) with the optimal constant $C_q = A_q^{1/q}$.

Theorem I.2. *Let Y_1, \dots, Y_n be jointly distributed nonnegative (non-independent) random variables, and let X_1, \dots, X_n be independent random variables such that each X_i has the same distribution as Y_i . Then, for every $q \geq 1$,*

$$\left\| \sum_{i=1}^n X_i \right\|_q \leq A_q^{1/q} \left\| \sum_{i=1}^n Y_i \right\|_q,$$

where A_q is the fractional Bell number which equals the q -th moment of the Poisson random variable with parameter 1 (see (1)). The constant $A_q^{1/q}$ is the optimal constant.

Proof: We first show that we cannot replace A_q with a smaller constant. Consider the following example. Let $Y_i^{(n)}, i \in [n]$ be random variables taking value 1 with probability $1/n$, and 0 with probability $1 - 1/n$. We generate $Y_i^{(n)}$'s as follows. We pick a random $j \in [n]$ and let $Y_j = 1$ and $Y_i = 0$ for $i \neq j$. Random variables $X_i^{(n)}$ are i.i.d. Bernoulli random variables with $\mathbb{E}[X_i^{(n)}] = 1/n$. Then, the sum $\sum_{i=1}^n Y_i^{(n)}$ always equals 1, and $\left\| \sum_{i=1}^n Y_i^{(n)} \right\|_q = 1$. As $n \rightarrow \infty$, the sum $\sum_{i=1}^n X_i^{(n)}$ converges in distribution to a Poisson random variable with parameter 1. Thus, $\left\| \sum_{i=1}^n X_i^{(n)} \right\|_q \rightarrow A_q^{1/q}$, and, hence, the constant $A_q^{1/q}$ cannot be improved.

Our proof is motivated by the example above. We prove the result for discrete finite random variables. The general case can be obtained using a standard argument by approximating random variables Y_1, \dots, Y_n with discrete random variables³.

Consider the finite probability space (Ω, \Pr) on which the random variables Y_1, \dots, Y_n are defined. Without loss of generality we may assume that $\Omega = \mathcal{Y}_1 \times \dots \times \mathcal{Y}_n$, where \mathcal{Y}_i is the range of the random variable Y_i . Then every elementary event is a vector $\omega = (y_1, \dots, y_n) \in \mathcal{Y}_1 \times \dots \times \mathcal{Y}_n$. The probability of the event $\omega = (y_1, \dots, y_n)$ equals

$$\Pr(\omega) = \Pr(Y_1 = y_1, \dots, Y_n = y_n).$$

Let $f_i(\omega)$ be the i -th coordinate of ω (or, more generally, the value the random variable Y_i takes when the elementary event $\omega \in \Omega$ occurs).

Let χ_ω be the indicator random variable of the event ω . Then, $\sum_{i=1}^n Y_i = \sum_{\omega \in \Omega} \sum_{i=1}^n \chi_\omega f_i(\omega)$. Each random variable χ_ω is a Bernoulli random variable with parameter $\Pr(\omega)$ i.e. $\Pr(\chi_\omega = 1) = \Pr(\omega)$. The random variables χ_ω are dependent: one and only one χ_ω equals 1 for any random outcome. As in the example above, we want to replace χ_ω 's

³In this paper, we only use the discrete version of this inequality.

with independent copies ψ_ω and, then, show that

$$\begin{aligned} \left\| \sum_{\omega \in \Omega} \psi_\omega \sum_{i=1}^n f_i(\omega) \right\|_q &\leq A_q^{1/q} \left\| \sum_{\omega \in \Omega} \chi_\omega \sum_{i=1}^n f_i(\omega) \right\|_q \\ &\equiv \left\| \sum_{i=1}^n Y_i \right\|_q. \end{aligned}$$

For technical reasons, we consider Poisson random variable instead of Bernoulli random variables. For each $\omega \in \Omega$, we define $n+1$ independent random variables $P_\omega^i, i \in [n]$ and P_ω on a new probability space Ω' . Each P_ω^i and P_ω is a Poisson random variable with parameter $\lambda_\omega = \Pr(\omega)$. Then, $\mathbb{E}[P_\omega^i] = \mathbb{E}[P_\omega] = \Pr(\omega)$.

We prove the following inequalities that imply the theorem.

1. $\left\| \sum_{i=1}^n X_i \right\|_q \leq \left\| \sum_{\omega \in \Omega} \sum_{i=1}^n f_i(\omega) P_\omega^i \right\|_q;$
2. $\left\| \sum_{\omega \in \Omega} \sum_{i=1}^n f_i(\omega) P_\omega^i \right\|_q \leq \left\| \sum_{\omega \in \Omega} \sum_{i=1}^n f_i(\omega) P_\omega \right\|_q;$
3. $\left\| \sum_{\omega \in \Omega} \sum_{i=1}^n f_i(\omega) P_\omega \right\|_q \leq A_q^{1/q} \left\| \sum_{i=1}^n Y_i \right\|_q.$

We split the proof into three main lemmas.

Lemma V.1. *Inequality 1 holds.*

Proof: We prove by induction on n the following inequality: for every $B \geq 0$,

$$\left\| B + \sum_{i=1}^n X_i \right\|_q \leq \left\| B + \sum_{\omega \in \Omega} \sum_{i=1}^n f_i(\omega) P_\omega^i \right\|_q.$$

For $n = 0$ this inequality trivially holds. For $n \geq 1$, we write

$$\begin{aligned} \left\| B + \sum_{i=1}^n X_i \right\|_q^q &= \mathbb{E} \left[\left(B + \sum_{i=1}^n X_i \right)^q \right] \\ &= \mathbb{E}_{X_n} \mathbb{E} \left[\left((B + X_n) + \sum_{i=1}^{n-1} X_i \right)^q \mid X_n \right] \\ &\leq \mathbb{E}_{X_n} \mathbb{E} \left[\left((B + X_n) + \sum_{i=1}^{n-1} \sum_{\omega \in \Omega} f_i(\omega) P_\omega^i \right)^q \mid X_n \right] \\ &= \mathbb{E} \left[\left((B + X_n) + \sum_{i=1}^{n-1} \sum_{\omega \in \Omega} f_i(\omega) P_\omega^i \right)^q \right]. \end{aligned}$$

Here, we used the inductive hypothesis with $B_* = B + X_n$. Let

$$S_{n-1} = \sum_{i=1}^{n-1} \sum_{\omega \in \Omega} f_i(\omega) P_\omega^i.$$

We need to show that

$$\mathbb{E} \left[\left(B + S_{n-1} + X_n \right)^q \right] \leq \mathbb{E} \left[\left(B + S_{n-1} + \sum_{\omega \in \Omega} f_n(\omega) P_\omega^n \right)^q \right].$$

We condition on S_{n-1} and prove that this inequality holds for every fixed value S'_{n-1} of the random variable S_{n-1} . Note that X_n and P_ω^n are independent from S_{n-1} . Let $B_\circ = B +$

S'_{n-1} . Since X_n and Y_n are identically distributed, we can replace X_n with Y_n . Thus, the inequality above follows from the inequality

$$\mathbb{E}\left[(B_\circ + Y_n)^q\right] \leq \mathbb{E}\left[(B_\circ + \sum_{\omega \in \Omega} f_n(\omega)P_\omega^n)^q\right].$$

Define a linear function $l : \mathbb{R}^\Omega \rightarrow \mathbb{R}$ and a non-linear function $g : \mathbb{R}^\Omega \rightarrow \mathbb{R}$ as follows: for $v \in \mathbb{R}^\Omega$ (the coordinates of v are indexed by $\omega \in \Omega$),

$$l(v) = B_\circ^q + \sum_{\omega \in \Omega} ((f_n(\omega) + B_\circ)^q - B_\circ^q)v_\omega;$$

$$g(v) = (B_\circ + \sum_{\omega \in \Omega} f_n(\omega)v_\omega)^q.$$

Note that if exactly one coordinate of v equals 1, and all other coordinates equal 0, then $l(v) = g(v)$. By Lemma V.2 (see below), if all coordinates of v are nonnegative integers then $g(v) \geq l(v)$. Let χ_ω be the indicator random variable of the elementary event ω ; and $\chi \in \mathbb{R}^\Omega$ be the vector with coordinates χ_ω . Observe that for any random outcome, exactly one coordinate of χ equals 1. Hence, $l(v) = g(v)$. Then,

$$(B_\circ + Y_n)^q \equiv (B_\circ + \sum_{\omega \in \Omega} f_n(\omega)\chi_\omega)^q = g(\chi) = l(\chi).$$

For $\chi = (\chi_\omega)$ and $P^n = (P_\omega^n)$, we have

$$\mathbb{E}\left[(B_\circ + X_n)^q\right] = \mathbb{E}\left[(B_\circ + Y_n)^q\right] = \mathbb{E}[l(\chi)] = l(\mathbb{E}[\chi]);$$

and

$$\mathbb{E}\left[(B_\circ + \sum_{\omega \in \Omega} f_n(\omega)P_\omega^n)^q\right] = \mathbb{E}[g(P^n)] \geq \mathbb{E}[l(P^n)]$$

$$= l(\mathbb{E}[P^n]).$$

However, $\mathbb{E}[\chi_\omega] = \Pr(\omega) = \mathbb{E}[P_\omega^n]$. Hence, $\mathbb{E}[\chi] = \mathbb{E}[P^n]$, and

$$\mathbb{E}\left[(B_\circ + X_n)^q\right] \leq \mathbb{E}\left[(B_\circ + \sum_{\omega \in \Omega} f_n(\omega)P_\omega^n)^q\right].$$

This finishes the proof. \blacksquare

Lemma V.2. *If all coordinates of v are nonnegative integers then $g(v) \geq l(v)$.*

Proof: Suppose that $g(v) \geq l(v)$ for some v . Fix a coordinate $\omega^* \in \Omega$, and let $\tilde{v}_\omega = v_\omega$ for $\omega \neq \omega^*$ and $\tilde{v}_{\omega^*} = v_{\omega^*} + 1$. That is, \tilde{v} alters from v only in the coordinate ω^* . We show that $g(\tilde{v}) \geq l(\tilde{v})$. Consider the function $h(t) = (t + f_n(\omega^*))^q - t^q$. This function is increasing for $t \geq 0$:

$$h'(t) = ((t + f_n(\omega^*))^q - t^q)'$$

$$= q((t + f_n(\omega^*))^{q-1} - t^{q-1}) > 0.$$

We have

$$g(\tilde{v}) - g(v) = h(B_\circ + \sum_{\omega \in \Omega} f_n(\omega)v_\omega) \geq h(B_\circ)$$

$$= (B_\circ + f_n(\omega^*))^q - B_\circ^q = l(\tilde{v}) - l(v).$$

Hence, $g(\tilde{v}) \geq l(\tilde{v})$.

For the zero vector $v = 0$, $g(v) = l(v)$. We can reach any nonnegative integer vector v , if we start from the zero vector and increment by 1 one coordinate at a time. Thus, $g(v) \geq l(v)$ for all nonnegative integer vectors v . \blacksquare

Lemma V.3. *Inequality 2 holds.*

We prove a slightly more general statement.

Lemma V.4. *Let P_j^i, P_j be independent nonnegative random variables (for $i \in \{1, \dots, n\}, j \in \{1, \dots, m\}$), let f_j^i be a sequence of nonnegative real numbers; and let $B \in \mathbb{R}^+$ be a nonnegative real number. Suppose that each P_j^i has the same distribution as P_j and $q \geq 1$. Then*

$$\mathbb{E}\left[\left(B + \sum_{j=1}^m \sum_{i=1}^n f_j^i P_j^i\right)^q\right] \leq \mathbb{E}\left[\left(B + \sum_{j=1}^m \sum_{i=1}^n f_j^i P_j\right)^q\right].$$

To get Lemma V.3 we apply Lemma V.4 with $f_j^i = f_i(\omega_j)$, $P_j^i = P_{\omega_j}^i$, $B = 0$, where $\omega_1, \dots, \omega_m$ is an arbitrary ordering of elementary events in Ω .

Proof: We prove Lemma V.4 by induction on m . For $m = 0$, the inequality trivially holds. Denote $F_j = \sum_{i=1}^n f_j^i$. We need to prove that

$$\mathbb{E}\left[\left(B + \sum_{j=1}^m \sum_{i=1}^n f_j^i P_j^i\right)^q\right] \leq \mathbb{E}\left[\left(B + \sum_{j=1}^m F_j P_j\right)^q\right].$$

Write

$$\mathbb{E}\left[\left(B + \sum_{j=1}^m \sum_{i=1}^n f_j^i P_j^i\right)^q\right] =$$

$$= \mathbb{E} \mathbb{E}\left[\left(\left(B + \sum_{i=1}^n f_m^i P_m^i\right) + \sum_{j=1}^{m-1} \sum_{i=1}^n f_j^i P_j^i\right)^q \mid P_m^1, \dots, P_m^n\right]$$

$$\leq \mathbb{E} \mathbb{E}\left[\left(\left(B + \sum_{i=1}^n f_m^i P_m^i\right) + \sum_{j=1}^{m-1} F_j P_j\right)^q \mid P_m^1, \dots, P_m^n\right]$$

$$= \mathbb{E}\left[\left(\left(B + \sum_{i=1}^n f_m^i P_m^i\right) + \sum_{j=1}^{m-1} F_j P_j\right)^q\right].$$

Here, we used the inductive hypothesis with $B_* = (B + \sum_{i=1}^n f_m^i P_m^i)$. Denote by B_\circ the random variable $B + \sum_{j=1}^{m-1} F_j P_j$. Then,

$$\mathbb{E}\left[\left(B + \sum_{j=1}^m \sum_{i=1}^n f_j^i P_j^i\right)^q\right] \leq \mathbb{E}\left[\left(B_\circ + \sum_{i=1}^n f_m^i P_m^i\right)^q\right].$$

Using convexity of the function $t \mapsto t^q$ for $q \geq 1$, we get

$$\left(B_\circ + \sum_{i=1}^n f_m^i P_m^i\right)^q = \left(\sum_{i=1}^n \frac{f_m^i}{F_m} (B_\circ + F_m P_m^i)\right)^q$$

$$\leq \sum_{i=1}^n \frac{f_m^i}{F_m} (B_\circ + F_m P_m^i)^q.$$

Each term $(B_\circ + F_m P_m^i)$ is distributed as $(B_\circ + F_m P_m)$, hence $\mathbb{E}[(B_\circ + F_m P_m^i)^q] = \mathbb{E}[(B_\circ + F_m P_m)^q]$, and

$$\begin{aligned} \mathbb{E}\left[\left(B + \sum_{j=1}^m \sum_{i=1}^n f_j^i P_j^i\right)^q\right] &\leq \mathbb{E}\left[\left(B_\circ + \sum_{i=1}^n f_m^i P_m\right)^q\right] \\ &\leq \sum_{i=1}^n \frac{f_m^i}{F_m} \mathbb{E}[(B_\circ + F_m P_m^i)^q] \\ &= \mathbb{E}[(B_\circ + F_m P_m)^q] \\ &= \mathbb{E}\left[\left(B + \sum_{j=1}^m F_j P_j\right)^q\right]. \end{aligned}$$

This concludes the proof. \blacksquare

Lemma V.5. *Inequality 3 holds.*

Proof: Let

$$P = \sum_{\omega \in \Omega} P_\omega.$$

The random variable P has the Poisson distribution with parameter 1, since random variables $P_\omega, \omega \in \Omega$, are independent and $\sum_{\omega} \mathbb{E}[P_\omega] = \sum_{\omega} \Pr(\omega) = 1$. We have

$$\begin{aligned} \left\| \sum_{\omega \in \Omega} \sum_{i=1}^n f_i(\omega) P_\omega \right\|_q^q &= \mathbb{E}\left[\left(\sum_{\omega \in \Omega} \sum_{i=1}^n f_i(\omega) P_\omega\right)^q\right] \quad (31) \\ &= \sum_{k=1}^{\infty} \mathbb{E}\left[\left(\sum_{\omega \in \Omega} \sum_{i=1}^n f_i(\omega) P_\omega\right)^q \mid P = k\right] \cdot \Pr(P = k) \\ &= \sum_{k=1}^{\infty} \mathbb{E}\left[\left(\sum_{\omega \in \Omega} \sum_{i=1}^n f_i(\omega) \frac{P_\omega}{k}\right)^q \mid P = k\right] \cdot (k^q \cdot \Pr(P = k)) \end{aligned}$$

Using convexity of the function $t \mapsto t^q$ for $q \geq 1$, we upper bound each term in the sum as follows:

$$\begin{aligned} \mathbb{E}\left[\left(\sum_{\omega \in \Omega} \sum_{i=1}^n f_i(\omega) \frac{P_\omega}{k}\right)^q \mid P = k\right] &\leq \\ &\leq \mathbb{E}\left[\sum_{\omega \in \Omega} \frac{P_\omega}{k} \left(\sum_{i=1}^n f_i(\omega)\right)^q \mid P = k\right] \\ &= \sum_{\omega \in \Omega} \mathbb{E}\left[\frac{P_\omega}{k} \mid P = k\right] \cdot \left(\sum_{i=1}^n f_i(\omega)\right)^q. \end{aligned}$$

We observe that $\mathbb{E}[P_\omega \mid P = k] = k \Pr(\omega)$, which follows from the following well known fact (see e.g., Feller [14], Section IX.9, Problem 6(b), p. 237).

Fact V.6. *Suppose P_1 and P_2 are independent Poisson random variables with parameters λ_1 and λ_2 . Then, for every $k \in \mathbb{N}$,*

$$\mathbb{E}[P_1 \mid P_1 + P_2 = k] = \frac{\lambda_1}{\lambda_1 + \lambda_2} k.$$

In our case, $P_1 = P_\omega$, $P_2 = \sum_{\omega' \neq \omega} P_{\omega'}$, $P_1 + P_2 = P$. Therefore, we have

$$\mathbb{E}\left[\left(\sum_{\omega \in \Omega} \sum_{i=1}^n f_i(\omega) \frac{P_\omega}{k}\right)^q \mid P = k\right] \leq \sum_{\omega \in \Omega} \Pr(\omega) \left(\sum_{i=1}^n f_i(\omega)\right)^q.$$

Plugging this inequality in (31), we obtain the desired bound

$$\begin{aligned} \left\| \sum_{\omega \in \Omega} \sum_{i=1}^n f_i(\omega) P_\omega \right\|_q^q &\leq \\ &\leq \sum_{\omega \in \Omega} \Pr(\omega) \left(\sum_{i=1}^n f_i(\omega)\right)^q \cdot \sum_{k=1}^{\infty} k^q \cdot \Pr(P = k) \\ &= \mathbb{E}\left[\left(\sum_{i=1}^n Y_i\right)^q\right] \cdot \mathbb{E}[P^q] = \left\| \sum_{i=1}^n Y_i \right\|_q^q \cdot \|P\|_q^q. \end{aligned}$$

\blacksquare

VI. NEGATIVELY ASSOCIATED RANDOM VARIABLES

The decoupling inequality (2) can be extended to *negatively associated* random variables X_1, \dots, X_n . The notion of negative association is defined as follows.

Definition VI.1 (Joag-Dev and Proschan [21]). *Random variables X_1, \dots, X_n are negatively associated if for all disjoint sets $I, J \subset [n]$ and all non-decreasing functions $f : \mathbb{R}^I \rightarrow \mathbb{R}$ and $g : \mathbb{R}^J \rightarrow \mathbb{R}$ the following inequality holds:*

$$\begin{aligned} \mathbb{E}[f(X_i, i \in I) \cdot g(X_j, j \in J)] &\leq \\ &\leq \mathbb{E}[f(X_i, i \in I)] \cdot \mathbb{E}[g(X_j, j \in J)]. \end{aligned}$$

Shao [31] showed that if X_1, \dots, X_n are negatively associated random variables, and X_1^*, \dots, X_n^* are independent random variables such that each X_i^* is distributed as X_i , then for every convex function $h : \mathbb{R} \rightarrow \mathbb{R}$,

$$\mathbb{E}[h(X_1 + \dots + X_n)] \leq \mathbb{E}[h(X_1^* + \dots + X_n^*)].$$

As an immediate corollary, for $h(x) = x^q$ (where $q \geq 1$), we have

$$\left\| \sum_{i=1}^n X_i \right\|_q \leq \left\| \sum_{i=1}^n X_i^* \right\|_q.$$

Therefore, the following corollary of Theorem I.2 holds.

Corollary VI.2. *Let Y_1, \dots, Y_n be jointly distributed nonnegative (non-independent) random variables, and let X_1, \dots, X_n be negatively associated random variables such that each X_i has the same distribution as Y_i . Then, for every $q \geq 1$,*

$$\left\| \sum_{i=1}^n X_i \right\|_q \leq A_q^{1/q} \left\| \sum_{i=1}^n Y_i \right\|_q,$$

where A_q is the fractional Bell number.

VII. GENERALIZATIONS

We can extend our results to a more general class of objective functions. Using our framework, we can solve combinatorial optimization problems with the objective function

$$\sum_{j \in [k]} f_j \left(\sum_{i \in [n]} d_{ij} y_i \right), \quad (32)$$

where f_j 's are arbitrary increasing *convex* functions satisfying $f_j(0) = 0$. In this case, the approximation ratio equals $A_{\{f\}} = \mathbb{E}[a_{\{f\}}(P_1)]$, where P_1 is a Poisson random variable

with parameter 1, and the function $a_{\{f\}}(t)$ is defined as $a_{\{f\}}(t) = \max\{f_j(tx)/f_j(x) : x > 0, j \in [k]\}$ for $t \in \mathbb{N}$. Note, that for $f(s) = cs^q$, $a_f(t) = t^q$ in Theorem I.2.

Similarly, we can solve maximization problems with the objective function (32) if f_j 's are arbitrary non-decreasing concave functions satisfying $f_j(0) = 0$. The approximation ratio equals $B_{\{f\}} = \mathbb{E}[b_{\{f\}}(P_1)]$, where $b_{\{f\}}(t) = \min\{f_j(tx)/f_j(x) : x > 0, j \in [k]\}$ for $t \in \mathbb{N}$. It is not hard to see that $B_{\{f\}} \geq (e-1)/e$. Indeed, in the worst case, $b_{\{f\}}(t) = 1$ for $t \geq 1$ and $b_{\{f\}}(t) = 0$ for $t = 0$, then $\mathbb{E}[b_{\{f\}}(P_1)] = \Pr(P_1 \geq 1) = 1 - 1/e$. This happens e.g., for the function $f(s) = \min\{s, 1\}$. Note that the approximation ratio of $(e-1)/e \approx 0.632$ for maximization problems of this form was previously known (see Calinescu et al. [7]). However, for some concave functions f we get a better approximation. For example, for $f(s) = \sqrt{s}$, we get an approximation ratio of $B_{\sqrt{s}} \approx 0.773$.

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APPENDIX A FIGURES

$q =$	1	1.25	1.5	1.75	2
$A_q =$	1	1.163	1.373	1.645	2

Fig. 1. The values of A_q for some $q \in [1, 2]$. All values are rounded up to three decimal places.

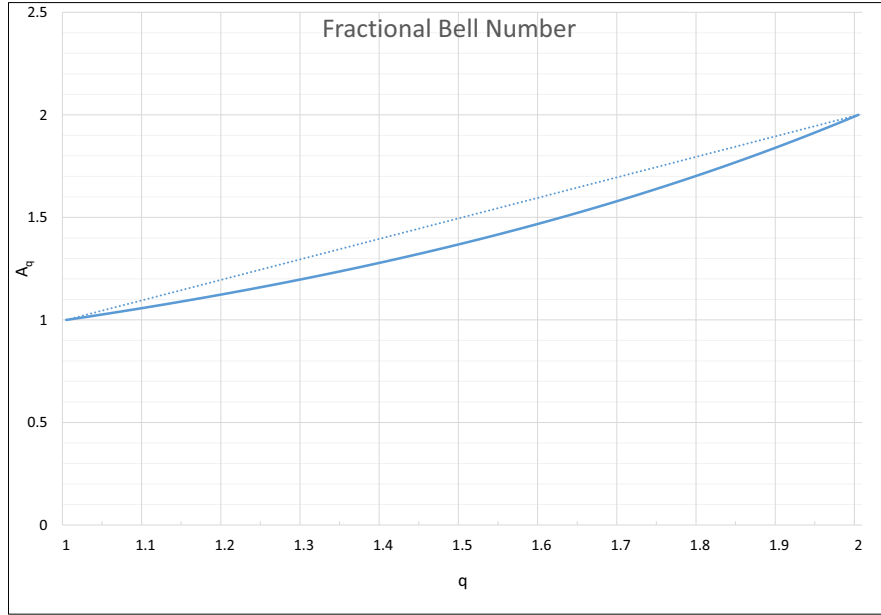


Fig. 2. Graph of A_q for $q \in [1, 2]$. Note that the function $q \mapsto A_q$ is convex; $A_1 = 1$ and $A_2 = 2$. Thus, $A_q \leq q$ for $q \in [1, 2]$.

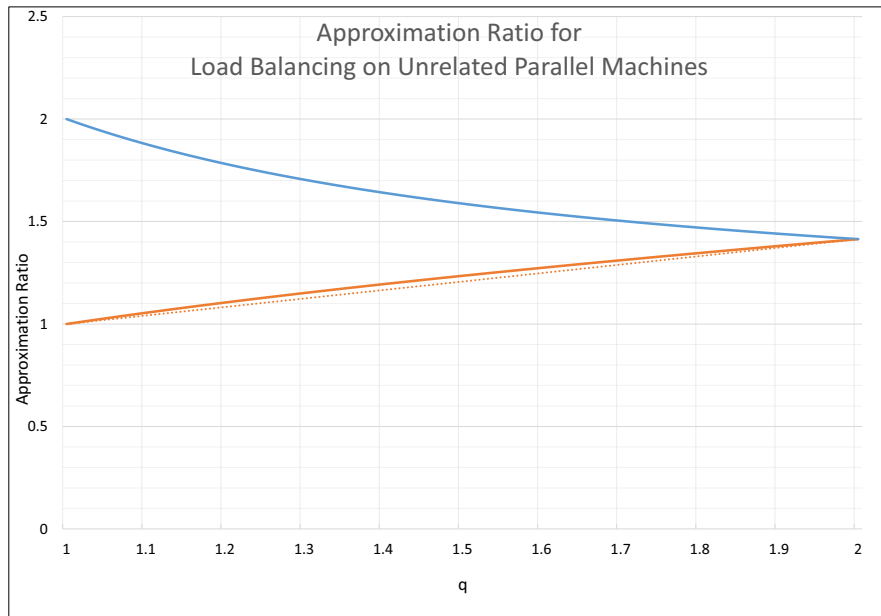


Fig. 3. Approximation factor of our algorithm for Load Balancing on Unrelated Parallel Machines – $A_q^{1/q}$ – is plotted in red (below). Approximation factor of the algorithm due to Kumar, Marathe, Parthasarathy and Srinivasan [22] – $2^{1/q}$ – is plotted in blue (above). The function $A_q^{1/q}$ can be well approximated by the linear function $1 + (\sqrt{2} - 1)(q - 1)$ in the interval $q \in [1, 2]$.