

## Learning Topic Models — Going beyond SVD

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**Abstract**—*Topic Modeling* is an approach used for automatic comprehension and classification of data in a variety of settings, and perhaps the canonical application is in uncovering thematic structure in a corpus of documents. A number of foundational works both in machine learning [17] and in theory [30] have suggested a probabilistic model for documents, whereby documents arise as a convex combination of (i.e. distribution on) a small number of *topic* vectors, each topic vector being a distribution on words (i.e. a vector of word-frequencies). Similar models have since been used in a variety of application areas; the *Latent Dirichlet Allocation* or LDA model of Blei et al. is especially popular.

Theoretical studies of topic modeling focus on learning the model’s parameters *assuming the data is actually generated from it*. Existing approaches for the most part rely on *Singular Value Decomposition* (SVD), and consequently have one of two limitations: these works need to either assume that each document contains only one topic, or else can only recover the *span* of the topic vectors instead of the topic vectors themselves.

This paper formally justifies *Nonnegative Matrix Factorization* (NMF) as a main tool in this context, which is an analog of SVD where all vectors are nonnegative. Using this tool we give the first polynomial-time algorithm for learning topic models without the above two limitations. The algorithm uses a fairly mild assumption about the underlying topic matrix called *separability*, which is usually found to hold in real-life data. Perhaps the most attractive feature of our algorithm is that it generalizes to yet more realistic models that incorporate topic-topic correlations, such as the *Correlated Topic Model* (CTM) and the *Pachinko Allocation Model* (PAM).

We hope that this paper will motivate further theoretical results that use NMF as a replacement for SVD – just as NMF has come to replace SVD in many applications.

### I. INTRODUCTION

Developing tools for automatic comprehension and classification of data — web pages, newspaper articles, images, genetic sequences, user ratings — is a holy grail of machine learning. *Topic Modeling* is an approach that has proved successful in all of the aforementioned settings, though for concreteness here we will focus on uncovering thematic structure of a corpus of documents (see e.g. [5], [7]).

In order to learn structure one has to posit the *existence* of structure, and in topic models one assumes a *generative model* for a collection of documents. Specifically, each document is represented as a vector of word-frequencies

(the *bag of words* representation). Seminal papers in theoretical CS (Papadimitriou et al. [30]) and machine learning (Hofmann’s *Probabilistic Latent Semantic Analysis* [17]) suggested that documents arise as a convex combination of (i.e. distribution on) a small number of *topic* vectors, where each topic vector is a distribution on words (i.e. a vector of word-frequencies). Each convex combination of topics thus is itself a distribution on words, and the document is assumed to be generated by drawing  $N$  independent samples from it. Subsequent work makes specific choices for the distribution used to generate topic combinations —the well-known *Latent Dirichlet Allocation* (LDA) model of Blei et al [7] hypothesizes a *Dirichlet* distribution (see Section IV).

Thus the topic modeling problem consists of fitting a good topic model to the document corpus. The prevailing approach in machine learning is to use local search (e.g. [12]) or other heuristics [35] in an attempt to find a *maximum likelihood* fit to the above model. For example, fitting to a corpus of newspaper articles may reveal fifty topic vectors corresponding to, say, politics, sports, weather, entertainment etc., and a particular article could be explained as a  $(1/2, 1/3, 1/6)$ -combination of the topics politics, sports, and entertainment. Unfortunately (and not surprisingly), the maximum likelihood estimation is  $NP$ -hard (see Section V) and consequently when using this paradigm, it seems necessary to rely on unproven heuristics even though these have well-known limitations (e.g. getting stuck in a local minima [12], [31]).

The work of Papadimitriou et al [30] (which also formalized the topic modeling problem) and a long line of subsequent work have attempted to give *provable* guarantees for the problem of learning the model parameters *assuming the data is actually generated from it*. This is in contrast to a maximum likelihood approach, which asks to find the closest-fit model for arbitrary data. The principal algorithmic problem is the following (see Section I-A for more details):

**Meta Problem in Topic Modeling:** *There is an unknown topic matrix  $A$  with nonnegative entries that is dimension  $n \times r$ , and a stochastically generated unknown matrix  $W$  that is dimension  $r \times m$ . Each column of  $AW$  is viewed as a probability distribution on rows, and for each column we are given  $N \ll n$  i.i.d. samples from the associated distribution.*

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**Goal:** Reconstruct  $A$  and parameters of the generating distribution for  $W$ .

The challenging aspect of this problem is that we wish to recover *nonnegative* matrices  $A, W$  with small inner-dimension  $r$ . The general problem of finding nonnegative factors  $A, W$  of specified dimensions when given the matrix  $AW$  (or a close approximation) is called the *Nonnegative Matrix Factorization* (NMF) problem (see [24], and [3] for a longer history) and it is NP-hard [34]. Lacking a tool to solve such problems, theoretical work has generally relied on the *Singular Value Decomposition* (SVD) which given the matrix  $AW$  will instead find factors  $U, V$  with both positive and negative entries. SVD has the feel of tool clustering — and its application in this setting seems to require assuming that each document has only *one* topic. In Papadimitriou et al [30] this is called the *pure documents* case and is solved under strong, additional assumptions about the topic matrix  $A$ . (See also [29] and the recent work of Anandkumar et al. [2] which completely solves this case using the method of moments.) Alternatively, other papers use SVD to recover the *span* of the columns of  $A$  (i.e. the topic vectors) [4], [22], [21], which suffices for some applications such as computing the inner product of two document vectors (in the space spanned by the topics) as a measure of their *similarity*.

These limitations of existing approaches —either restricting to one topic per document, or else learning only the span of the topics instead of the topics themselves—are quite serious. In practice documents are much more faithfully described as a distribution on topics and indeed for a wide range of applications one needs the actual topics and not just their span — such as when browsing a collection of documents without a particular query phrase in mind, or when tracking how topics evolve over time (see [5] for a survey of various applications). Here we consider what we believe to be a much weaker assumption — *separability*. Indeed, this property has already been identified as a natural one in the machine learning community [13] and has been empirically observed to hold in topic matrices fitted to various types of data [6].

Separability requires that each topic has some near-perfect indicator word — a word that we call the *anchor word* for this topic — that appears with reasonable probability in that topic but with negligible probability in all other topics (e.g., “401k” could be an anchor word for the topic “personal finance”). We give a formal definition in Section I-A. This property is particularly natural in the context of topic modeling, where the number of distinct words (dictionary size) is very large compared to the number of topics. In a typical application, it is common to have a dictionary size in the thousands or tens of thousands, but the number of topics is usually somewhere in the range from fifty to a hundred. Note that separability does *not* mean that an anchor word always occurs —in fact, a typical document may be very likely to contain *no* anchor words. Instead, separability

says that when an anchor word does occur, this is a strong indicator that the corresponding topic is in the mixture used to generate the document.

Recently, we gave a polynomial time algorithm to solve NMF under the condition that the topic matrix  $A$  is separable [3]. The intuition that underlies this algorithm is that the set of anchor words can be thought of as extreme points (in a geometric sense) of the dictionary. This condition can be used to identify all of the anchor words and then also the nonnegative factors. Ideas from this algorithm are a key ingredient in our present paper, but our focus is on the question:

**Question.** *What if we are not given the true matrix  $AW$ , but are instead given a few samples (say, a hundred samples) from the distribution represented by each column?*

The main technical challenge in adapting our earlier NMF algorithm is that each document vector is a *very poor* approximation to the corresponding column of  $AW$  — it is *too noisy* in any reasonable measure of noise. Nevertheless, the core insights of our NMF algorithm still apply. Note that it is impossible to learn the matrix  $W$  to within arbitrary accuracy. (Indeed, this is information theoretically impossible even if we knew the topic matrix  $A$  as well as the distribution from which the columns of  $W$  are generated.) So we *cannot* in general give an estimator that converges to the true matrix  $W$ , and yet we *can* give an estimator that converges to the true topic matrix  $A$ ! (For an overview of our algorithm, see the first paragraph of Section III.)

We hope that this application of our NMF algorithm is just a starting point and other theoretical results can use NMF as a replacement for SVD — just as NMF has come to replace SVD in several applied settings. In addition, the geometric problems that underly NMF are not yet fully understood and there are many interesting theoretical challenges that remain.

**Practical Issues.** The estimates of runtimes throughout the paper are possibly too pessimistic. As mentioned in the conclusions section, simple variations of the algorithms in this paper run very fast —much more so than existing software for topic models.

#### A. Our Results

Here we formally define the topic modeling (learning) problem which we informally introduced above. There is an unknown *topic matrix*  $A$  which is of dimension  $n \times r$  (i.e.  $n$  is the dictionary size) and each column of  $A$  is a distribution on  $[n]$ . There is an unknown  $r \times m$  matrix  $W$  each of whose columns is itself a distribution (i.e. a convex combination) on  $[r]$ . The columns of  $W$  are i.i.d. samples from a distribution  $\mathcal{T}$  which belongs to a known family, e.g., Dirichlet distributions, but whose parameters are unknown. Thus each column of  $AW$  being a convex combination of distributions is itself a distribution on  $[n]$ , and the algorithm’s

input consists of  $N$  i.i.d. samples for each column of  $AW$ . Here  $N$  is the document size and is assumed to be a constant for simplicity. Our algorithm can be easily adapted to work when the documents have different sizes.

The algorithm’s running time will necessarily depend upon various model parameters, since distinguishing a very small parameter from zero imposes a lower bound on the number of samples needed. The first such parameter is a quantitative version of *separability*, which was presented above as a natural assumption in context of topic modeling.

**Definition I.1** ( $p$ -Separable Topic Matrix). An  $n \times r$  matrix  $A$  is  $p$ -separable if for each  $i$  there is some row  $\pi(i)$  of  $A$  that has a single nonzero entry which is in the  $i^{\text{th}}$  column and it is at least  $p$ .

The next parameter measures the lowest probability with which a topic occurs in the distribution that generates columns of  $W$ .

**Definition I.2** (Topic Imbalance). The *topic imbalance* of the model is the ratio between the largest and smallest expected entries in a column of  $W$ , in other words,  $a = \max_{i,j \in [r]} \frac{\mathbf{E}[X_i]}{\mathbf{E}[X_j]}$  where  $X \in \mathbb{R}^r$  is a random weighting of topics chosen from the distribution.

Finally, we require that topics stay identifiable despite sampling-induced noise. To formalize this, we define a matrix that will be important throughout this paper:

**Definition I.3** (Topic-Topic Covariance Matrix  $R(\mathcal{T})$ ). If  $\mathcal{T}$  is the distribution that generates the columns of  $W$ , then  $R(\mathcal{T})$  is defined as an  $r \times r$  matrix whose  $(i, j)$ th entry is  $E[X_i X_j]$  where  $X_1, X_2, \dots, X_r$  is a vector chosen from  $\mathcal{T}$ .

Let  $\gamma > 0$  be a lower bound on the  $\ell_1$ -condition number of the matrix  $R(\mathcal{T})$ . This is defined in Section II, but for a  $r \times r$  matrix it is within a factor of  $\sqrt{r}$  of the smallest singular value. Our algorithm will work for any  $\gamma$ , but the number of documents we require will depend (polynomially) on  $1/\gamma$ :

**Theorem I.4** (Main). *There is a polynomial time algorithm that learns the parameters of a topic model if the number of documents is at least*

$$m \geq \max \left\{ O \left( \frac{\log n \cdot a^4 r^6}{\epsilon^2 p^6 \gamma^2 N} \right), O \left( \frac{\log r \cdot a^2 r^4}{\gamma^2} \right) \right\},$$

where the three parameters  $a, p, \gamma$  are as defined above. The algorithm learns the topic-term matrix  $A$  up to additive error  $\epsilon$ . Moreover, when the number of documents is also larger than  $O \left( \frac{\log r \cdot r^2}{\epsilon^2} \right)$  the algorithm can learn the topic-topic covariance matrix  $R(\mathcal{T})$  up to additive error  $\epsilon$ .

As noted earlier, we are able to recover the topic matrix even though we do not always recover the parameters of the column distribution  $\mathcal{T}$ . In some special cases we can also recover the parameters of  $\mathcal{T}$ , e.g. when this distribution is

Dirichlet, as happens in the popular *Latent Dirichlet Allocation* (LDA) model [7], [5]. This is done in Section IV-A by computing a lower bound on the  $\gamma$  in terms of the parameter for the Dirichlet distribution, which allows us with some other ideas (see Section IV-B) to recover the parameters of  $\mathcal{T}$  from the co-variance matrix  $R(\mathcal{T})$ .

Recently the basic LDA model has been refined to allow correlation among different topics, which is more realistic. See for example the *Correlated Topic Model* (CTM) [8] and the *Pachinko Allocation Model* (PAM) [26]. Perhaps the most attractive aspect of our algorithm is that it extends to these models as well: we can learn the topic matrix, even though we cannot always identify  $\mathcal{T}$ . In real data, there are *always* topics that are closely correlated (or very anti-correlated) and we believe that this extra generality is the reason our algorithm returns high-quality topics on real data.

*Comparison with related works:* (i) We rely crucially on separability. But prior works assume a single topic per document, which can be thought of as a stronger “separability” assumption about  $W$  instead of  $A$ . (ii) After posting a draft of this paper, a subsequent paper by Anandkumar et.al. [1] gave an algorithm to recover parameters of an LDA model without requiring  $A$  to be separable. These results are incomparable since we require separability but can allow topic correlations. We believe that allowing topic correlations is crucial when working with real data (and have found empirical evidence that supports this conclusion). (iii) We remark that some prior approaches learn the span of  $A$  instead of  $A$  require large document sizes (on the order of the number of words in the dictionary!). *By contrast we can work with documents of length 2.*

## II. TOOLS FOR (NOISY) NONNEGATIVE MATRIX FACTORIZATION

### A. Various Condition Numbers

Central to our arguments will be various notions of matrices being “far” from being low-rank. The most interesting one for our purposes was introduced by Kleinberg and Sandler [21] in the context of collaborative filtering and can be thought of as an  $\ell_1$ -analogue to the smallest singular value of a matrix.

**Definition II.1** ( $\ell_1$  Condition Number). If a matrix  $B$  has nonnegative entries and all rows sum to one then its  $\ell_1$  Condition Number  $\Gamma(B)$  is defined as:

$$\Gamma(B) = \min_{\|x\|_1=1} \|xB\|_1.$$

If  $B$  does not have row sums of one then  $\Gamma(B)$  is equal to  $\Gamma(DB)$  where  $D$  is the diagonal matrix such that  $DB$  has row sums of one.

For example, if the rows of  $B$  have disjoint support then  $\Gamma(B) = 1$  and in general the quantity  $\Gamma(B)$  can be thought of a measure of how close two distributions on *disjoint* sets

of rows can be. Note that if  $x$  is an  $n$ -dimensional real vector,  $\|x\|_2 \leq \|x\|_1 \leq \sqrt{n}\|x\|_2$  and hence (if  $\sigma_{\min}(B)$  is the smallest singular value of  $B$ ) we have:

$$\frac{1}{\sqrt{n}}\sigma_{\min}(B) \leq \Gamma(B) \leq \sqrt{m}\sigma_{\min}(B).$$

The above notions of condition number will be most relevant in the context of the topic-topic covariance matrix  $R(\mathcal{T})$ . We shall always use  $\gamma$  to denote the  $\ell_1$  condition number of  $R(\mathcal{T})$ . The definition of condition number will be preserved even when we estimate the topic-topic covariance matrix using random samples.

**Lemma II.2.** *When  $m > 5 \log r / \epsilon_0^2$ , with high probability the matrix  $R = \frac{1}{m}WW^T$  is entry-wise close to  $R(\mathcal{T})$  with error  $\epsilon_0$ . Further, when  $\epsilon_0 < \gamma/4ar^2$  where  $a$  is topic imbalance, the matrix  $R$  has  $\ell_1$  condition number at least  $\gamma/2$ .*

*Proof:* Since  $\mathbf{E}[W_i W_i^T] = R(\mathcal{T})$ , the first part of the lemma follows by a Chernoff bound and a union bound. The second part follows because  $R(\mathcal{T})$  has  $\ell_1$  condition number  $\gamma$ , and for unit vector  $v$  the vector  $vR$  can change by at most  $ar \cdot r\epsilon_0$  in  $\ell_1$  norm. The extra factor  $ar$  comes from the normalization of rows of  $R$ . ■

In our previous work on nonnegative matrix factorization [3] we defined a different measure of “distance” from singular which is essential to the polynomial time algorithm for NMF:

**Definition II.3** ( $\beta$ -robustly simplicial). If each column of a matrix  $A$  has unit  $\ell_1$  norm, then we say it is  $\beta$ -robustly simplicial if no column in  $A$  has  $\ell_1$  distance smaller than  $\beta$  to the convex hull of the remaining columns in  $A$ .

The following claim clarifies the interrelationships of these latter condition numbers.

**Claim II.4.** (i) If  $A$  is  $p$ -separable then  $A^T$  has  $\ell_1$  condition number at least  $p$ . (ii) If  $A^T$  has all row sums equal to 1 then  $A$  is  $\beta$ -robustly simplicial for  $\beta = \Gamma(A^T)/2$ .

We shall see that the  $\ell_1$  condition number for product of matrices is at least the product of the  $\ell_1$  condition numbers. The main application of this composition is to show that the matrix  $R(\mathcal{T})A^T$  (or the empirical version  $RA^T$ ) is at least  $\Omega(\gamma p)$ -robustly simplicial. The following lemma will play a crucial role in analyzing our main algorithm:

**Lemma II.5** (Composition Lemma). *If  $B$  and  $C$  are matrices with  $\ell_1$  condition number  $\Gamma(B) \geq \gamma$  and  $\Gamma(C) \geq \beta$ , then  $\Gamma(BC)$  is at least  $\beta\gamma$ . Specifically, when  $A$  is  $p$ -separable the matrix  $R(\mathcal{T})A^T$  is at least  $\gamma p/2$ -robustly simplicial.*

*Proof:* For any vector  $x$  we have  $\|xBC\|_1 \geq \Gamma(C)\|xB\|_1 \geq \Gamma(C)\Gamma(B)\|x\|_1$ . For the matrix  $R(\mathcal{T})A^T$ , by Claim II.4 we know the matrix  $A^T$  has  $\ell_1$  condition number at least  $p$ . Hence  $\Gamma(R(\mathcal{T})A^T)$  is at least  $\gamma p$  and

again by Claim II.4 the matrix is  $\gamma p/2$ -robustly simplicial. ■

### B. Noisy NMF under Separability

A key ingredient is an approximate NMF algorithm from [3] which can recover an approximate nonnegative matrix factorization  $\tilde{M} \approx AW$  when the  $\ell_1$  distance between each row of  $\tilde{M}$  and the corresponding row in  $AW$  is small. We emphasize that this is not enough for our purposes, since the term-by-document matrix  $\tilde{M}$  will have a substantial amount of noise (when compared to its expectation) precisely because the number of words in a document  $N$  is much smaller than the dictionary size  $n$ . Rather, we will apply to the Gram matrix  $\tilde{M}\tilde{M}^T$  the algorithms given in the following theorem and its improvement in the subsequent theorem.

**Theorem II.6** (Robust NMF Algorithm [3]). *Suppose  $M = AW$  where  $W$  and  $M$  are normalized to have rows sum up to 1,  $A$  is separable and  $W$  is  $\gamma$ -robustly simplicial. Let  $\epsilon = O(\gamma^2)$ . There is a polynomial time algorithm that given  $\tilde{M}$  such that for all rows  $\|\tilde{M}^i - M^i\|_1 < \epsilon$ , finds a  $W'$  such that  $\|W'^i - W^i\|_1 < 10\epsilon/\gamma + 7\epsilon$ . Further every row  $W'^i$  in  $W'$  is a row in  $M$ . The corresponding row in  $M$  can be represented as  $(1 - O(\epsilon/\gamma^2))W^i + O(\epsilon/\gamma^2)W^{-i}$ . Here  $W^{-i}$  is a vector in the convex hull of other rows in  $W$  with unit length in  $\ell_1$  norm.*

In this paper we have an incomparable goal than in [3]. Our goal is not to recover estimates to the anchor words that are close in  $\ell_1$ -norm but rather to recover almost anchor words (word whose row in  $A$  has almost all its weight on a single coordinate). Hence, we will be able to achieve better bounds by treating this problem directly, and we give a substitute for the above theorem. The proof of the Theorem can be found in the full version.

**Theorem II.7.** *Suppose  $M = AW$  where  $W$  and  $M$  are normalized to have rows sum up to 1,  $A$  is separable and  $W$  is  $\gamma$ -robustly simplicial. When  $\epsilon < \gamma/100$  there is a polynomial time algorithm that given  $\tilde{M}$  such that for all rows  $\|\tilde{M}^i - M^i\|_1 < \epsilon$ , finds  $r$  row (almost anchor words) in  $\tilde{M}$ . The  $i$ -th almost anchor word corresponds to a row in  $M$  that can be represented as  $(1 - O(\epsilon/\gamma))W^i + O(\epsilon/\gamma)W^{-i}$ . Here  $W^{-i}$  is a vector in the convex hull of other rows in  $W$  with unit length in  $\ell_1$  norm.*

## III. ALGORITHM FOR LEARNING A TOPIC MODEL: PROOF OF THEOREM I.4

First it is important to understand why separability helps in nonnegative matrix factorization and the exact role played by the anchor words. Suppose the NMF algorithm is given a matrix  $AB$ . If  $A$  is  $p$ -separable then this means that  $A$  contains a diagonal matrix (up to row permutations). Thus a scaled copy of each row of  $B$  is present as a row in  $AB$ . In fact, if we knew the anchor words of  $A$ , then by looking

at the corresponding rows of  $AB$  we could “read off” the corresponding row of  $B$  (up to scaling) and use these in turn to recover all of  $A$ . Thus the anchor words constitute the “key” that “unlocks” the factorization, and indeed the main step of our earlier NMF algorithm was a geometric procedure to identify the anchor words. When one is given a noisy version of  $AB$  the analogous notion is “almost anchor” words which correspond to rows of  $AB$  that are “very close” to rows of  $B$ ; see Theorem II.7.

Next we sketch how to apply these insights to learning topic models. Let  $M$  denote the given term-by-document matrix, each of whose columns describes the empirical word frequencies in the documents. It is obtained from sampling  $AW$  and thus is an extremely noisy approximation to  $AW$ . Our algorithm starts by forming the Gram matrix  $MM^T$ , which can be thought of as an empirical word-word covariance matrix. In fact as the number of documents increases  $\frac{1}{m}MM^T$  tends to a limit  $Q = \frac{1}{m}E[AWW^TA]$ , implying  $Q = AR(\mathcal{J})A^T$ . (See Lemma III.7.) Imagine that we are given the exact matrix  $Q$  instead of a noisy approximation. Notice that  $Q$  is a product of *three* nonnegative matrices, the first of which is  $p$ -separable and the last is the transpose of the first. NMF at first glance seems too weak to help find such factorizations. However, if we think of  $Q$  as a product of *two* nonnegative matrices,  $A$  and  $R(\mathcal{J})A^T$ , then our NMF algorithm [3] can at least identify the anchor words of  $A$ . As noted above, these suffice to recover  $R(\mathcal{J})A^T$ , and then (using the anchor words of  $A$  again) all of  $A$  as well. See Section III-A for details.

The complication is that we are not given  $Q$  but merely a good approximation to it. Now our NMF algorithm allows us to recover “almost anchor” words of  $A$ , and the crux of the proof is Section III-B showing that these suffice to recover provably good estimates to  $A$  and  $WW^T$ . This uses (mostly) bounds from matrix perturbation theory and inter-relationships of condition numbers mentioned in Section II.

For simplicity we assume the following condition on the topic model, which we will see in Section III-D can be assumed without loss of generality:

(\*) *The number of words,  $n$ , is at most  $4ar/\epsilon$ .*

Please see Algorithm 1: Main Algorithm for description of the algorithm. Note that  $R$  is our shorthand for  $\frac{1}{m}WW^T$ , which as noted converges to  $R(\mathcal{J})$  as the number of documents increases.

#### A. Recover $R$ and $A$ with Anchor Words

We first describe how the recovery procedure works in an “idealized” setting (Algorithm 2:RECOVER WITH TRUE ANCHOR WORDS), when we are given the exact value of  $ARA^T$  and a set of anchor words – one for each topic. We can permute the rows of  $A$  so that the anchor words are exactly the first  $r$  words. Therefore  $A^T = (D, U^T)$  where  $D$  is a diagonal matrix. Note that  $D$  is not necessarily the identity matrix (nor even a scaled copy of the identity

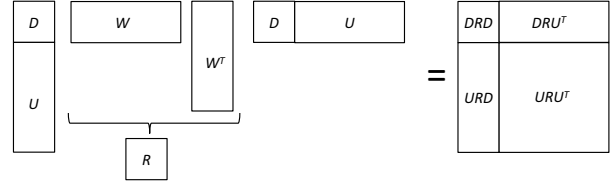


Figure 1. The matrix  $Q$

matrix), but we do know that the diagonal entries are at least  $p$ . We apply the same permutation to the rows and columns of  $Q$ . As illustrated in Figure 1 the submatrix formed by the first  $r$  rows and  $r$  columns is exactly  $DRD$ . Similarly, the submatrix consisting of the first  $r$  rows is exactly  $DRA^T$ . We can use these two matrices to compute  $R$  and  $A$ , in this idealized setting (and we will use the same basic strategy in the general case, but need only be more careful about how the errors accumulate in our algorithm).

Our algorithm has exact knowledge of the matrices  $DRD$  and  $DRA^T$  and so the main task is to recover the diagonal matrix  $D$ . Given  $D$ , we can then compute  $A$  and  $R$  (for the Dirichlet Allocation we can also compute its parameters - i.e. the  $\vec{\alpha}$  so that  $R(\alpha) = R$ ). The key idea to this algorithm is that the row sums of  $DR$  and  $DRA^T$  are the same, and we can use the row sums of  $DR$  to set up a system of linear constraints on the diagonal entries of  $D^{-1}$ .

**Lemma III.1.** *When the matrix  $Q$  is exactly equal to  $ARA^T$  and we know the set of anchor words, RECOVER WITH TRUE ANCHOR WORDS outputs  $A$  and  $R$  correctly.*

*Proof:* The Lemma is straight forward from Figure 1 and the procedure. By Figure 1 we can find the exact value of  $DRA^T$  and  $DRD$  in the matrix  $Q$ . Step 2 of recover computes  $DR\vec{1}$  by computing  $DRA^T\vec{1}$ . The two vectors are equal because  $A$  is the topic-term matrix and its columns sum up to 1, in particular  $A^T\vec{1} = \vec{1}$ .

In Step 3, since  $R$  is invertible by Lemma II.2,  $D$  is a diagonal matrix with entries at least  $p$ , the matrix  $DRD$  is also invertible. Therefore there is a unique solution  $\vec{z} = (DRD)^{-1}DR\vec{1} = D^{-1}\vec{1}$ . Also  $D\vec{z} = \vec{1}$  and hence  $DDiag(z) = I$ . Finally, using the fact that  $DDiag(z) = I$ , the output in step 4 is just  $(DR)^{-1}DRA^T = A^T$ , and the output in step 5 is equal to  $R$ . ■

#### B. Recover $R$ and $A$ with Almost Anchor Words

What if we are not given the exact anchor words, but are given words that are “close” to anchor words? In general we cannot hope to recover the true anchor words, but nevertheless a good approximation to these will be enough to recover  $R$  and  $A$ .

When we restrict  $A$  to the rows corresponding to “almost” anchor words, the submatrix will not be diagonal. However,

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**Algorithm 1.** MAIN ALGORITHM, **Output:**  $R$  and  $A$ 

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1) Query the oracle for  $m$  documents, where

$$m = \max \left\{ O \left( \frac{\log n \cdot a^4 r^6}{\epsilon^2 p^6 \gamma^2 N} \right), O \left( \frac{\log r \cdot a^2 r^4}{\gamma^2} \right), O \left( \frac{\log r \cdot r^2}{\epsilon^2} \right) \right\}$$

- 2) Split the words of each document into two halves, and let  $\tilde{M}$ ,  $\tilde{M}'$  be the term-by-document matrix with first and second half of words respectively.
- 3) Compute word-by-word matrix  $Q = \frac{4}{N^2 m} \tilde{M} \tilde{M}'^T$
- 4) Apply the “Robust NMF” algorithm of Theorem II.7 to  $Q$  which returns  $r$  words that are “almost” the anchor words of  $A$ .
- 5) Use these  $r$  words as input to RECOVER WITH ALMOST ANCHOR WORDS to compute  $R = \frac{1}{m} W W^T$  and  $A$
- 

**Algorithm 2.** RECOVER WITH TRUE ANCHOR WORDS**Input:**  $r$  anchor words, **Output:**  $R$  and  $A$ 

---

- 1) Permute the rows and columns of  $Q$  so that the anchor words appear in the first  $r$  rows and columns
- 2) Compute  $DRA^T \vec{1}$  (which is equal to  $DR \vec{1}$ )
- 3) Solve for  $\vec{z}$ :  $DRD\vec{z} = DR \vec{1}$ .
- 4) Output  $A^T = ((DRD \text{Diag}(z))^{-1} DRA^T)$ .
- 5) Output  $R = (\text{Diag}(z) DRD \text{Diag}(z))$ .
- 

it will be close to a diagonal in the sense that the submatrix will be a diagonal matrix  $D$  multiplied by  $E$ , where  $E$  is entry-wise close to the identity matrix (and the diagonal entries of  $D$  are at least  $\Omega(p)$ ). Here we analyze the same procedure as above and show that it still recovers  $A$  and  $R$  (approximately) even when given “almost” anchor words instead of true anchor words. For clarity we state the procedure again in Algorithm 3: RECOVER WITH ALMOST ANCHOR WORDS. The guarantees at each step are different than before, but the implementation of the procedure is the same. Notice that here we permute the rows of  $A$  (and hence the rows and columns of  $Q$ ) so that the “almost” anchor words returned by Theorem II.6 appear first and the submatrix  $A$  on these rows is equal to  $DE$ .

Here, we still assume that the matrix  $Q$  is exactly equal to  $ARA^T$  and hence the first  $r$  rows of  $Q$  form the submatrix  $DERA^T$  and the first  $r$  rows and columns are  $DERE^T D$ . The complication here is that  $\text{Diag}(z)$  is not necessarily equal to  $D^{-1}$ , since the matrix  $E$  is not necessarily the identity. However, we can show that  $\text{Diag}(z)$  is “close” to  $D^{-1}$  if  $E$  is suitably close to the identity matrix – i.e. given good enough proxies for the anchor words, we can bound the error of the above recovery procedure. We write  $E = I + Z$ . Intuitively when  $Z$  has only small entries  $E$  should behave like the identity matrix. In particular,  $E^{-1}$  should have only small off-diagonal entries. We make this precise through the following lemmas:

**Lemma III.2.** *Let  $E = I + Z$  and  $\sum_{i,j} |Z_{i,j}| = \epsilon < 1/2$ , then  $E^{-1} \vec{1}$  is a vector with entries in the range  $[1 - 2\epsilon, 1 + 2\epsilon]$ .*

*Proof:*  $E$  is clearly invertible because the spectral norm of  $Z$  is at most  $1/2$ . Let  $\vec{b} = E^{-1} \vec{1}$ . Since  $E = I + Z$  we multiply  $E$  on both sides to get  $\vec{b} + Z\vec{b} = \vec{1}$ . Let  $b_{max}$

be the largest absolute value of any entry of  $b$  ( $b_{max} = \max |b_i|$ ). Consider the entry  $i$  where  $b_{max}$  is achieved, we know  $b_{max} = |b_i| \leq 1 + |(Zb)_i| \leq 1 + \sum_j |Z_{i,j}| |b_j| \leq 1 + \epsilon b_{max}$ . Thus  $b_{max} \leq 1/(1 - \epsilon) \leq 2$ . Now all the entries in  $Z\vec{b}$  are within  $2\epsilon$  in absolute value, and we know that  $\vec{b} = \vec{1} + Z\vec{b}$ . Hence all the entries of  $b$  are in the range  $[1 - 2\epsilon, 1 + 2\epsilon]$ , as desired. ■

**Lemma III.3.** *Let  $E = I + Z$  and  $\sum_{i,j} |Z_{i,j}| = \epsilon < 1/2$ , then the columns of  $E^{-1} - I$  have  $\ell_1$  norm at most  $2\epsilon$ .*

*Proof:* Without loss of generality, we can consider just the first column of  $E^{-1} - I$ , which is equal to  $(E^{-1} - I)\vec{e}_1$ , where  $\vec{e}_1$  is the indicator vector that is one on the first coordinate and zero elsewhere.

The approach is similar to that in Lemma III.2. Let  $\vec{b} = (E^{-1} - I)\vec{e}_1$ . Left multiply by  $E = (I + Z)$  and we obtain  $\vec{b} + Z\vec{b} = -Z\vec{e}_1$ . Hence  $\vec{b} = -Z(\vec{b} + \vec{e}_1)$ . Let  $b_{max}$  be the largest absolute value of entries of  $\vec{b}$  ( $b_{max} = \max |b_i|$ ). Let  $i$  be the entry in which  $b_{max}$  is achieved. Then

$$b_{max} = |b_i| \leq |(Z\vec{b})_i| + |(Z\vec{e}_1)_i| \leq \epsilon b_{max} + \epsilon$$

Therefore  $b_{max} \leq \epsilon/(1 - \epsilon) \leq 2\epsilon$ . Further, the  $\|\vec{b}\|_1 \leq \|Z\vec{e}_1\|_1 + \|Z\vec{b}\|_1 \leq \epsilon + 2\epsilon^2 \leq 2\epsilon$ . ■

Now we are ready to show that the procedure RECOVER WITH ALMOST ANCHOR WORDS succeeds when given “almost” anchor words:

**Lemma III.4.** *When the matrix  $Q$  is exactly equal to  $ARA^T$ , the matrix  $A$  restricted to almost anchor words is  $DE$  where  $E - I$  has  $\ell_1$  norm  $\epsilon < 1/10$  when viewed as a vector, procedure RECOVER WITH ALMOST ANCHOR WORDS outputs  $A$  such that each column of  $A$  has  $\ell_1$  error at most  $6\epsilon$ . The matrix  $R$  has additive error  $Z_R$  whose  $\ell_1$  norm when viewed as a vector is at most  $8\epsilon$ .*

---

**Algorithm 3.** RECOVER WITH ALMOST ANCHOR WORDS

**Input:**  $r$  "almost" anchor words, **Output:**  $R$  and  $A$ 


---

- 1) Permute the rows and columns of  $Q$  so that the "almost" anchor words appear in the first  $r$  rows and columns.
  - 2) Compute  $DERA^T \vec{1}$  (which is equal to  $DER \vec{1}$ )
  - 3) Solve for  $\vec{z}$ :  $DERE^T D \vec{z} = DER \vec{1}$ .
  - 4) Output  $A^T = ((DERE^T D \text{Diag}(z))^{-1} DERA^T)$ .
  - 5) Output  $R = (\text{Diag}(z) DERE^T D \text{Diag}(z))$ .
- 

*Proof:* Since  $Q$  is exactly  $ARA^T$ , our algorithm is given  $DERA^T$  and  $DERE^T D$  with no error. In Step 3, since  $D$ ,  $E$  and  $R$  are all invertible, we have

$$\vec{z} = (DERE^T D)^{-1} DER \vec{1} = D^{-1} (E^T)^{-1} \vec{1}$$

Ideally we would want  $\text{Diag}(z) = D^{-1}$ , and indeed  $D \text{Diag}(z) = \text{Diag}((E^T)^{-1} \vec{1})$ . From Lemma III.2, the vector  $(E^T)^{-1} \vec{1}$  has entries in the range  $[1 - 2\epsilon, 1 + 2\epsilon]$ , thus each entry of  $\text{Diag}(z)$  is within a  $(1 \pm 2\epsilon)$  multiplicative factor from the corresponding entry in  $D^{-1}$ .

Consider the output in Step 4. Since  $D$ ,  $E$ ,  $R$  are invertible, the first output is

$$(DERE^T D \text{Diag}(z))^{-1} DERA^T = (D \text{Diag}(z))^{-1} (E^T)^{-1} A^T$$

Our goal is to bound the  $\ell_1$  error of the columns of the output compared to the corresponding columns of  $A$ . Notice that it is sufficient to show that the  $j^{\text{th}}$  row of  $(D \text{Diag}(z))^{-1} (E^T)^{-1}$  is close (in  $\ell_1$  distance) to the indicator vector  $\vec{e}_j^T$ .

**Claim III.5.** For each  $j$ ,  $\|\vec{e}_j^T (D \text{Diag}(z))^{-1} (E^T)^{-1} - \vec{e}_j^T\|_1 \leq 5\epsilon$

*Proof:* Again, without loss of generality we can consider just the first row. From Lemma III.3  $\vec{e}_1^T (E^T)^{-1}$  has  $\ell_1$  distance at most  $2\epsilon$  to  $\vec{e}_1^T$ .  $(D \text{Diag}(z))^{-1}$  has entries in the range  $[1 - 3\epsilon, 1 + 3\epsilon]$ . And so

$$\begin{aligned} & \|\vec{e}_1^T (D \text{Diag}(z))^{-1} (E^T)^{-1} - \vec{e}_1^T\|_1 \\ & \leq \|\vec{e}_1^T (D \text{Diag}(z))^{-1} (E^T)^{-1} - \vec{e}_1^T (E^T)^{-1}\|_1 \\ & \quad + \|\vec{e}_1^T (E^T)^{-1} - \vec{e}_1^T\|_1 \end{aligned}$$

The last term can be bounded by  $2\epsilon$ . Consider the first term on the right hand side: The vector  $\vec{e}_1^T (D \text{Diag}(z))^{-1} - \vec{e}_1^T$  has one non-zero entry (the first one) whose absolute value is at most  $3\epsilon$ . Hence, from Lemma III.3 the first term can be bounded by  $6\epsilon^2 \leq 3\epsilon$ , and this implies the claim. ■

The first row of  $(D \text{Diag}(z))^{-1} (E^T)^{-1} A^T$  is  $A_1 + z^T A$  where  $z$  is a vector with  $\ell_1$  norm at most  $5\epsilon$ . So every column of  $A$  is recovered with  $\ell_1$  error at most  $6\epsilon$ .

Consider the second output of the algorithm. The output is  $\text{Diag}(z) DERE^T D \text{Diag}(z)$  and we can write  $\text{Diag}(z) D = I + Z_1$  and  $E = I + Z_2$ . The leading error are  $Z_1 R + Z_2 R + R Z_1 + R Z_2$  and hence the  $\ell_1$  norm of the leading error term (when treated as a vector) is at most  $6\epsilon$  and other terms are

of order  $\epsilon^2$  and can safely be bounded by  $2\epsilon$  for suitably small  $\epsilon$ . ■

Finally we consider the general case (in which there is additive noise in Step 1): we are not given  $ARA^T$  exactly. We are given  $Q$  which is close to  $ARA^T$  (by Lemma III.7). We will bound the accumulation of this last type of error. Suppose in Step 1 of RECOVER we obtain  $DERA^T + U$  and  $DERE^T D + V$  and furthermore the entries of  $U$  and  $U \vec{1}$  have absolute value at most  $\epsilon_1$  and the matrix  $V$  has  $\ell_1$  norm  $\epsilon_2$  when viewed as a vector.

**Lemma III.6.** If  $\epsilon, \epsilon_1, \epsilon_2$  are sufficiently small, RECOVER outputs  $A$  such that each entry of  $A$  has additive error at most  $O(\epsilon + (ra\epsilon_2/p^3 + \epsilon_1 r/p^2)/\gamma)$ . Also the matrix  $R$  has additive error  $Z_R$  whose  $\ell_1$  norm when viewed as a vector is at most  $O(\epsilon + (ra\epsilon_2/p^3 + \epsilon_1 r/p^2)/\gamma)$ .

The main idea of the proof is to write  $DERE^T D + V$  as  $DER(E^T + V')D$ . In this way the error  $V$  can be translated to an error  $V'$  on  $E$  and Lemma III.4 can be applied.

*Proof:* We shall follow the proof of Lemma III.4. First can express the error term  $V$  instead as  $V = (DER)V'(D)$ . This is always possible because all of  $D$ ,  $E$ ,  $R$  are invertible. Moreover, the  $\ell_1$  norm of  $V'$  when viewed as a vector is at most  $8ra\epsilon_2/\gamma p^3$ , because this norm will grow by a factor of at most  $1/p$  when multiplied by  $D^{-1}$ , a factor of at most 2 when multiplied by  $E^{-1}$  and at most  $ra/\Gamma(R)$  when multiplied by  $R^{-1}$ . The bound of  $\Gamma(R)$  comes from Lemma II.2, we lose an extra  $ra$  because  $R$  may not have rows sum up to 1.

Hence  $DERE^T D + V = DER(E^T + V')D$  and the additive error for  $DERE^T D$  can be transformed into error in  $E$ , and we will be able to apply the analysis in Lemma III.4.

Similarly, we can express the error term  $U$  as  $U = DERU'$ . Entries of  $U'$  have absolute value at most  $8\epsilon_1 r/\gamma p^2$ . The right hand side of the equation in step 3 is equal to  $DER \vec{1} + U \vec{1}$  so the error is at most  $\epsilon_1$  per entry. Following the proof of Lemma III.4, we know  $\text{Diag}(z)D$  has diagonal entries within  $1 \pm (2\epsilon + 16\epsilon_2/\gamma p^3 + 2\epsilon_1)$ .

Now we consider the output. The output for  $A^T$  is equal to  $(DER(E^T + V')D \text{Diag}(z))^{-1} DER(A^T + U')$ , which is  $(D \text{Diag}(z))^{-1} (E^T + V')^{-1} (A^T + U')$ . Here we know  $(E^T + V')^{-1} - I$  has  $\ell_1$  norm at most  $O(\epsilon + ra\epsilon_2/\gamma p^3)$  per row,  $(D \text{Diag}(z))$  is a diagonal matrix with entries in  $1 \pm O(\epsilon + ra\epsilon_2/\gamma p^3 + \epsilon_1)$ , entries of  $U'$  has absolute value  $O(\epsilon_1 r/\gamma p^2)$ . Following the proof of Lemma III.4 the final entry-wise error

of  $A$  is roughly the sum of these three errors, and is bounded by  $O(\epsilon + (ra\epsilon_2/p^3 + \epsilon_1 r/p^2)/\gamma)$  (Notice that Lemma III.4 gives bound for  $\ell_1$  norm of rows, which is stronger. Here we switched to entry-wise error because the entries of  $U$  are bounded while the  $\ell_1$  norm of  $U$  might be large).

Similarly, the output of  $R$  is equal to  $\text{Diag}(z)(DERE^T D + V)\text{Diag}(z)$ . Again we write  $\text{Diag}(z)D = I + Z_1$  and  $E = I + Z_2$ . The extra term  $\text{Diag}(z)V\text{Diag}(z)$  is small because the entries of  $z$  are at most  $2/p$  (otherwise  $\text{Diag}(z)D$  won't be close to identity). The error can be bounded by  $O(\epsilon + (ra\epsilon_2/p^3 + \epsilon_1 r/p^2)/\gamma)$ . ■

Now in order to prove our main theorem we just need to show that when number of documents is large enough, the matrix  $Q$  is close to the  $ARA^T$ , and plug the error bounds into Lemma III.6. We state the convergence of  $Q$  below and defer the details to the full version.

### C. Error Bounds for $Q$

Here we state the error bound for matrix  $Q$ , whose proof we defer to the full version.

**Lemma III.7.** *When  $m > \frac{50 \log n}{N\epsilon_Q^2}$ , with high probability all entries of  $Q - \frac{1}{m}AWW^T A^T$  have absolute value at most  $\epsilon_Q$ . Further, the  $\ell_1$  norm of rows of  $Q$  are also  $\epsilon_Q$  close to the  $\ell_1$  norm of the corresponding row in  $\frac{1}{m}AWW^T A^T$ .*

### D. Reducing Dictionary Size

So far we have assumed that the number of distinct words is not too large. Here, we give a simple gadget to demonstrate that this is true without loss of generality:

**Lemma III.8.** *The general case can be reduced to an instance in which there are at most  $4ar/\epsilon$  words all of which (with at most one exception) occur with probability at least  $\epsilon/4ar$ .*

The proof is straightforward and the idea is to collect all words that occur infrequently and “merge” all of these words into a aggregate word that we will call the *runoff word*. We defer the proof to the full version.

## IV. THE DIRICHLET SUBCASE

Here we demonstrate that the parameters of a Dirichlet distribution can be (robustly) recovered from just the covariance matrix  $R(\mathcal{T})$ . Hence an immediate corollary is that our main learning algorithm can recover both the topic matrix  $A$  and the distribution that generates columns of  $W$  in a *Latent Dirichlet Allocation* (LDA) Model [7], provided that  $A$  is separable. We believe that this algorithm may be of practical use, and provides the first alternative to local search and (unproven) approximation procedures for this inference problem [35], [12], [7].

The Dirichlet distribution is parametrized by a vector  $\alpha$  of positive reals and is a natural family of continuous

multivariate probability distributions. The support of the Dirichlet Distribution is the unit simplex whose dimension is the same as the dimension of  $\alpha$ . Let  $\alpha$  be a  $r$  dimensional vector. Then for a vector  $\theta \in \mathbb{R}^r$  in the  $r$  dimensional simplex, its probability density is given by  $Pr[\theta|\alpha] = \frac{\Gamma(\sum_{i=1}^r \alpha_i)}{\prod_{i=1}^r \Gamma(\alpha_i)} \prod_{i=1}^r \theta_i^{\alpha_i-1}$ , where  $\Gamma$  is the Gamma function. In particular, when all of the  $\alpha_i$ 's are equal to one, the Dirichlet Distribution is just the uniform distribution on the probability simplex.

The expectation and variance of  $\theta_i$ 's are easy to compute given the parameters  $\alpha$ . We denote  $\alpha_0 = \|\alpha\|_1 = \sum_{i=1}^r \alpha_i$ , then the ratio  $\alpha_i/\alpha_0$  should be interpreted as the “size” of the  $i$ -th variable  $\theta_i$ , and  $\alpha_0$  controls whether the distribution is concentrated in the interior (when  $\alpha_0$  is large) or near the boundary (when  $\alpha_0$  is small). The first two moments of the Dirichlet distribution are:  $\mathbf{E}[\theta_i] = \frac{\alpha_i}{\alpha_0}$ ,

$$\mathbf{E}[\theta_i \theta_j] = \begin{cases} \frac{\alpha_i \alpha_j}{\alpha_0(\alpha_0+1)} & \text{when } i \neq j \\ \frac{\alpha_i(\alpha_i+1)}{\alpha_0(\alpha_0+1)} & \text{when } i = j \end{cases}$$

Suppose the Dirichlet distribution has  $\max \alpha_i / \min \alpha_i = a$  and the sum of parameters is  $\alpha_0$ ; we give an algorithm that computes close estimates to the vector of parameters  $\alpha$  given a sufficiently close estimate to the co-variance matrix  $R(\mathcal{T})$  (Theorem IV.3). Combining this with Theorem I.4, we obtain the following corollary:

**Theorem IV.1.** *There is an algorithm that learns the topic matrix  $A$  with high probability up to an additive error of  $\epsilon$  from at most*

$$m = \max \left\{ O \left( \frac{\log n \cdot a^6 r^8 (\alpha_0 + 1)^4}{\epsilon^2 p^6 N} \right), O \left( \frac{\log r \cdot a^2 r^4 (\alpha_0 + 1)^2}{\epsilon^2} \right) \right\}$$

*documents sampled from the LDA model and runs in time polynomial in  $n, m$ . Furthermore, we recover the parameters of the Dirichlet distribution to within an additive  $\epsilon$ .*

Our main goal in this section is to bound the  $\ell_1$ -condition number of the Dirichlet distribution (see Section IV-A), and using this we show how to recover the parameters of the distribution from its covariance matrix (Section IV-B).

### A. Condition Number of a Dirichlet Distribution

There is a well-known meta-principle that if a matrix  $W$  is chosen by picking its columns independently from a fairly diffuse distribution, then it will be far from low rank. However, our analysis will require us to prove an explicit lower bound on  $\Gamma(R(\mathcal{T}))$ . We now prove such a bound when the columns of  $W$  are chosen from a Dirichlet distribution with parameter vector  $\alpha$ . We note that it is easy to establish such bounds for other types of distributions as well. Recall that we defined  $R(\mathcal{T})$  in Section I, and here we will abuse notation and throughout this section we will denote by  $R(\alpha)$



---

**Algorithm 4.** DIRICHLET( $R$ ), **Input:**  $R$ , **Output:**  $\alpha$  (vector of parameters)

---

- 1) Set  $\alpha/\alpha_0 = R\bar{1}$ .
  - 2) Let  $i$  be the row with smallest  $\ell_1$  norm, let  $u = R_{i,i}$  and  $v = \alpha_i/\alpha_0$ .
  - 3) Set  $\alpha_0 = \frac{1-u/v}{u/v-v}$ .
  - 4) Output  $\alpha = \alpha_0 \cdot (\alpha/\alpha_0)$ .
- 

the matrix  $R(\mathcal{J})$  where  $\mathcal{J}$  is the Dirichlet distribution with parameter  $\alpha$ .

Let  $\alpha_0 = \sum_{i=1}^r \alpha_i$ . The mean, variance and co-variance for a Dirichlet distribution are well-known, from which we observe that  $R(\alpha)_{i,j}$  is equal to  $\frac{\alpha_i \alpha_j}{\alpha_0(\alpha_0+1)}$  when  $i \neq j$  and is equal to  $\frac{\alpha_i(\alpha_i+1)}{\alpha_0(\alpha_0+1)}$  when  $i = j$ .

**Lemma IV.2.** *The  $\ell_1$  condition number of  $R(\alpha)$  is at least  $\frac{1}{2(\alpha_0+1)}$ .*

*Proof:* As the entries  $R(\alpha)_{i,j}$  is  $\frac{\alpha_i \alpha_j}{\alpha_0(\alpha_0+1)}$  when  $i \neq j$  and  $\frac{\alpha_i(\alpha_i+1)}{\alpha_0(\alpha_0+1)}$  when  $i = j$ , after normalization  $R(\alpha)$  is just the matrix  $D' = \frac{1}{\alpha_0+1} (\alpha \times (1, 1, \dots, 1) + I)$  where  $\times$  is outer product and  $I$  is the identity matrix.

Let  $x$  be a vector such that  $|x|_1 = 1$  and  $|D'x|_1$  achieves the minimum in  $\Gamma(R(\alpha))$  and let  $I = \{i | x_i \geq 0\}$  and let  $J = \bar{I}$  be the complement. We can assume without loss of generality that  $\sum_{i \in I} x_i \geq |\sum_{i \in J} x_i|$  (otherwise just take  $-x$  instead). The product  $D'x$  is  $\sum_{i \in I} \frac{x_i}{\alpha_0+1} \alpha + \frac{1}{\alpha_0+1} x$ . The first term is a nonnegative vector and hence for each  $i \in I$ ,  $(D'x)^i \geq 0$ . This implies that  $|D'x|_1 \geq \frac{1}{\alpha_0+1} \sum_{i \in I} x_i \geq \frac{1}{2(\alpha_0+1)}$ . ■

### B. Recovering the Parameters of a Dirichlet Distribution

When the covariance matrix  $R(\alpha)$  is recovered with error  $\epsilon_R$  in  $\ell_1$  norm when viewed as a vector, we can use Algorithm 4: DIRICHLET to compute the vector  $\alpha$ .

**Theorem IV.3.** *When the covariance matrix  $R(\alpha)$  is recovered with error  $\epsilon_R$  in  $\ell_1$  norm when viewed as a vector, the procedure DIRICHLET( $R$ ) learns the parameter of the Dirichlet distribution with error at most  $O(ar(\alpha_0 + 1)\epsilon_R)$ .*

*Proof:* The  $\alpha_i/\alpha_0$ 's all have error at most  $\epsilon_R$ . The value  $u$  is  $\frac{\alpha_i}{\alpha_0} \frac{\alpha_i+1}{\alpha_0+1} \pm \epsilon_R$  and the value  $v$  is  $\alpha_i/\alpha_0 \pm \epsilon_R$ . Since  $v \geq 1/ar$  we know the error for  $u/v$  is at most  $2ar\epsilon_R$ . Finally we need to bound the denominator  $\frac{\alpha_i+1}{\alpha_0+1} - \frac{\alpha_i}{\alpha_0} > \frac{1}{2(\alpha_0+1)}$  (since  $\frac{\alpha_i}{\alpha_0} \leq 1/r \leq 1/2$ ). Thus the final error is at most  $5ar(\alpha_0 + 1)\epsilon_R$ . ■

## V. MAXIMUM LIKELIHOOD ESTIMATION IS HARD

Here we observe that computing the Maximum Likelihood Estimate (MLE) of the parameters of a topic model is  $NP$ -hard. We call this problem the Topic Model Maximum Likelihood Estimation (TM-MLE) problem:

**Definition V.1** (TM-MLE). Given  $m$  documents and a target of  $r$  topics, the TM-MLE problem asks to compute the topic

matrix  $A$  that has the largest probability of generating the observed documents (when the columns of  $W$  are generated by a uniform Dirichlet distribution).

Surprisingly, this appears to be the first proof that computing the MLE estimate in a topic model is indeed computationally hard, although its hardness is certainly to be expected. On a related note, Sontag and Roy [32] recently proved that *given the topic matrix* and a document, computing the Maximum A Posteriori (MAP) estimate for the distribution on topics that generated this document is  $NP$ -hard. Here we will establish that TM-MLE is  $NP$ -hard via a reduction from the MIN-BISECTION problem: In MIN-BISECTION the input is a graph with  $n$  vertices ( $n$  is an even integer), and the goal is to partition the vertices into two equal sized sets of  $n/2$  vertices each so as to minimize the number of edges crossing the cut.

**Theorem V.2.** *There is a polynomial time reduction from MIN-BISECTION to TM-MLE ( $r = 2$ ).*

We defer the proof to the full version. We remark that the canonical solutions in our reduction are all *separable*, and hence this reduction applies even when the topic matrix  $A$  is known (and required) to be separable. So, even in the case of a separable topic matrix, it is  $NP$ -hard to compute the MLE. Yet, here we have given an efficient estimator that converges to the true (separable) topic matrix  $A$  when the data is actually generated according to the LDA model.

## VI. CONCLUSIONS

Though the goal of the paper is design of an algorithm with theoretical guarantees, the actual algorithm turns out to be practical. A straightforward implementation (using a more efficient, LP-free subroutine to find anchor words but no other tuning) runs much faster than state-of-the-art software for topic models, and gives results of comparable quality. For example on the UCI ‘‘Bag of Words’’ dataset with New York Times articles [14] we fit 200 topics in only 10 minutes on a dataset with 300000 articles with a vocabulary of size 102660, whereas MALLET [28] takes several hours. A detailed study of its performance is underway and will be reported soon. The separability assumption seems benign on such datasets. In fact our machine learning colleagues suggest that real-life topic matrices satisfy even stronger *separability* assumptions, e.g., the presence of *many* anchor words per topic instead of a single one. Leveraging this promising suggestion, is an open problem.

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